

1 **Supplementary Information:**

2 **Neuromorphic High-Throughput Imaging via** 3 **Transport-Induced Differential Sensing**

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14 **Supplementary Note 1: Throughput-constrained spatiotemporal sampling**

15 We consider sensing over a spatiotemporal domain $\Omega \times T$ under a fixed throughput constraint. Let $w(\mathbf{x}, t) \geq 0$
 16 denote a spatiotemporal sampling density such that the total number of statistically independent measurements
 17 satisfies

$$\int_{\Omega} \int_T w(\mathbf{x}, t) dt d\mathbf{x} = B, \quad (\text{S1})$$

18 where B denotes the available sensing throughput.

19 For synchronous frame-based acquisition, the sampling density is signal-independent and separable,

$$w_F(\mathbf{x}, t) = w_F(\mathbf{x}) w_F(t). \quad (\text{S2})$$

20 Here $w_F(\mathbf{x})$ represents spatial sampling on the sensor lattice and is proportional to the number of pixels N_{pix} ,
 21 which defines the native spatial resolution, while $w_F(t)$ represents temporal sampling and is proportional to the
 22 frame rate f_{fps} . Substituting Equation (S2) into Equation (S1) yields

$$B \propto N_{\text{pix}} f_{\text{fps}}, \quad (\text{S3})$$

23 which explicitly captures the coupling between spatial resolution and temporal sampling imposed by a fixed
 24 throughput budget in synchronous imaging.

25 In contrast, asynchronous event-based sensing induces a signal-dependent sampling density. Under locally
 26 rigid sweep motion and approximately stationary texture statistics within $\Omega \times T$, a separable approximation
 27 remains valid,

$$w_E(\mathbf{x}, t) \approx w_E(\mathbf{x}) w_E(t). \quad (\text{S4})$$

28 The spatial activation density $w_E(\mathbf{x})$ is governed by temporal variations of the logarithmic intensity. Under
 29 brightness constancy, these variations are determined by the motion-aligned first-order increment $-\nabla L(\mathbf{x}) \cdot \mathbf{v}$,
 30 leading to the activation model

$$w_E(\mathbf{x}) \propto \mathbf{1}(|-\nabla L(\mathbf{x}) \cdot \mathbf{v}| > c), \quad (\text{S5})$$

31 where c denotes the fixed temporal contrast threshold of the event sensor. This defines an active spatial subset

$$\Omega' = \{\mathbf{x} \in \Omega : |-\nabla L(\mathbf{x}) \cdot \mathbf{v}| > c\}, \quad \Omega' \subseteq \Omega, \quad (\text{S6})$$

32 and therefore $\#\Omega' \leq \#\Omega$ on the discrete sensor lattice.

33 Under the same global throughput constraint Equation (S1), asynchronous sensing reallocates measurements
 34 from uniform coverage of Ω to the activated subset Ω' , which is enriched with motion-aligned gradient infor-
 35 mation. The temporal factor $w_E(t)$ corresponds to the event triggering rate and remains approximately stable
 36 under stationary texture statistics; a probabilistic justification is provided in a later supplementary note. This
 37 reallocation explains how, under a fixed throughput budget, asynchronous sensing concentrates measurements on
 38 informative first-order structures rather than uniformly sampling zeroth-order photometric content.

39 Figure S1 provides a schematic illustration of how a fixed spatio-temporal bandwidth budget constrains typical
40 imaging system designs, positioning frame-based cameras along a trade-off between spatial sampling density
41 and temporal throughput, and highlighting how asynchronous event sensing enables an alternative operating
42 regime through hybrid fusion. Within this perspective, conventional frame cameras appear in two limiting forms:
43 high-resolution multi-megapixel CMOS sensors that prioritize dense spatial sampling at the expense of temporal
44 throughput, and single-pixel or few-pixel computational imaging systems based on compressive sampling, which
45 trade spatial parallelism for ultrafast temporal acquisition¹.

Supplementary Note 2: Circuit-origin of event measurements and Physical operator

We briefly summarize the circuit-level origin of event measurements and their physical interpretation, which motivates the linear event generation operator used in the main text.

As illustrated in Figure S2, an event camera pixel converts incident photocurrent into a logarithmic voltage signal through a logarithmic $I - V$ circuit. Let $I(\mathbf{x}, t)$ denote the incident optical intensity and $L(\mathbf{x}, t) \triangleq \log I(\mathbf{x}, t)$ the corresponding logarithmic intensity. A pixel continuously monitors the temporal change of $L(\mathbf{x}, t)$ and emits an event whenever the accumulated increment exceeds a fixed contrast threshold c determined by the comparator circuit,

$$\Delta L(\mathbf{x}_k, t_k) = L(\mathbf{x}_k, t_k + \Delta t_k) - L(\mathbf{x}_k, t_k) = p_k c, \quad p_k \in \{-1, +1\}, \quad (\text{S7})$$

where p_k denotes the polarity of the log-intensity change and Δt_k is the inter-event interval at pixel \mathbf{x}_k .

The threshold c is a hardware-defined parameter governed by the comparator and reset circuitry, and is treated as constant throughout this work. Upon threshold crossing, the pixel voltage is reset, resulting in a sequence of polarity-coded events that encode the sign of the temporal log-intensity increment, as schematically shown in Figure S2b.

Under the photometric constancy assumption for rigid motion, the temporal derivative of the logarithmic intensity satisfies

$$\frac{\partial L(\mathbf{x}, t)}{\partial t} = -\nabla L(\mathbf{x}, t) \cdot \mathbf{v}(t), \quad (\text{S8})$$

where $\mathbf{v}(t)$ denotes the instantaneous image-plane velocity. Over a sufficiently short temporal neighborhood, ∇L and \mathbf{v} can be treated as locally constant, yielding a first-order finite-difference approximation

$$\Delta L(\mathbf{x}; \Delta t) \approx -\nabla L(\mathbf{x}) \cdot \mathbf{v} \Delta t. \quad (\text{S9})$$

Combining the circuit-triggering condition in Equation (S7) with the motion-induced increment in Equation (S9), the event stream provides multiple measurements of a motion-aligned first-order feature of the logarithmic intensity field. After motion compensation to a reference time, these measurements can be aggregated to form a linear observation of the form

$$\mathbf{b}_E = \mathbf{A}_\mathbf{v} L(\mathbf{x}) + \mathbf{n}_E, \quad \mathbf{A}_\mathbf{v} \triangleq -\mathbf{v} \cdot \nabla, \quad (\text{S10})$$

where $L(\mathbf{x}) = \log I(\mathbf{x})$, $\mathbf{A}_\mathbf{v}$ denotes the event generation operator induced by motion, and \mathbf{n}_E captures the effective noise after temporal aggregation.

This formulation establishes a direct link between the circuit-level event triggering mechanism and the linear feature-domain measurement operator used in the main text.

72 **Supplementary Note 3: Probabilistic modeling of the event camera**

73 This note formalizes how the motion-aligned first-order increment

$$\mu(\mathbf{x}) \triangleq -\nabla L(\mathbf{x}) \cdot \mathbf{v}, \quad (\text{S11})$$

74 is converted by the event circuit into discrete stochastic events, and how temporal accumulation yields a
75 continuous-valued statistic that is approximately linear in μ and can be used as a differentiable measurement.

76 **S3.1 Instantaneous polarity as a thresholded random variable**

77 Following Supplementary Note 2, the event circuit compares a noisy log-domain increment against a symmetric
78 threshold $\pm c$. Specifically, within a sufficiently short interval δt , define the (pre-threshold) triggering variable

$$Z \sim \mathcal{N}(\mu, \sigma_Z^2), \quad (\text{S12})$$

79 where μ is the motion-aligned log-intensity increment in Equation (S11) and σ_Z^2 summarizes the effective
80 uncertainty induced by photon statistics and electronic readout (Supplementary Note 2). The ternary polarity
81 outcome in that interval is

$$S = \begin{cases} +1, & Z > c, \\ -1, & Z < -c, \\ 0, & |Z| \leq c. \end{cases} \quad (\text{S13})$$

82 Therefore, the probabilities of positive and negative events are

$$p_+(\mu) = \Pr(S = +1) = 1 - \Phi\left(\frac{c - \mu}{\sigma_Z}\right), \quad p_-(\mu) = \Pr(S = -1) = 1 - \Phi\left(\frac{c + \mu}{\sigma_Z}\right), \quad (\text{S14})$$

83 and the expected polarity (the mean signed response) is

$$m(\mu) \triangleq \mathbb{E}[S] = p_+(\mu) - p_-(\mu) = \Phi\left(\frac{\mu + c}{\sigma_Z}\right) - \Phi\left(\frac{c - \mu}{\sigma_Z}\right), \quad (\text{S15})$$

84 which is a symmetric sigmoidal function of μ (see Supplementary Fig. SS3).

85 **S3.2 Temporal accumulation and Skellam-to-Gaussian approximation**

86 Consider a temporal window of duration Δt consisting of $N = \Delta t / \delta t$ independent trials of S . Let

$$N_+ = \#\{S_i = +1\}, \quad N_- = \#\{S_i = -1\}, \quad N_0 = \#\{S_i = 0\}, \quad (\text{S16})$$

87 then

$$(N_+, N_-, N_0) \sim \text{Multinomial}(N; p_+(\mu), p_-(\mu), 1 - p_+(\mu) - p_-(\mu)). \quad (\text{S17})$$

88 When δt is sufficiently small, the accumulated positive and negative counts are well approximated by two
89 independent Poisson processes,

$$N_+ \sim \text{Poisson}(\Lambda_+), \quad N_- \sim \text{Poisson}(\Lambda_-), \quad \Lambda_{\pm} \triangleq r_{\pm}(\mu) \Delta t, \quad r_{\pm}(\mu) \triangleq \frac{p_{\pm}(\mu)}{\delta t}. \quad (\text{S18})$$

90 Define the accumulated bipolar count (signed accumulation)

$$C \triangleq N_+ - N_- \tag{S19}$$

91 Under Equation (S18), C follows a Skellam distribution with moments

$$\mathbb{E}[C] = \Lambda_+ - \Lambda_-, \quad \text{Var}(C) = \Lambda_+ + \Lambda_- \tag{S20}$$

92 We then define the time-averaged polarity statistic

$$\hat{g} \triangleq \frac{C}{\Delta t} \tag{S21}$$

93 It follows that

$$\mathbb{E}[\hat{g}] = r_+(\mu) - r_-(\mu), \quad \text{Var}(\hat{g}) = \frac{r_+(\mu) + r_-(\mu)}{\Delta t} \tag{S22}$$

94 Moreover, for moderate-to-large $\Lambda_+ + \Lambda_-$, the Skellam distribution approaches a Gaussian by the central limit
95 effect, yielding the approximation

$$\hat{g} \approx \mathcal{N}\left(r_+(\mu) - r_-(\mu), \frac{r_+(\mu) + r_-(\mu)}{\Delta t}\right) \tag{S23}$$

96 This provides a continuous-valued, approximately Gaussian measurement whose variance decreases as Δt
97 increases, i.e., temporal accumulation suppresses stochastic fluctuations.

98 **S3.3 Near-linear response around $\mu = 0$**

99 In the small-increment regime relevant to sub-pixel motion, $m(\mu)$ in Equation (S15) admits a local linearization
100 around $\mu = 0$,

$$m(\mu) \approx \eta \mu, \quad \eta \triangleq m'(0) = \frac{2}{\sigma_Z} \phi\left(\frac{c}{\sigma_Z}\right) \tag{S24}$$

101 where $\phi(\cdot)$ is the standard Gaussian probability density function. Since $r_{\pm}(\mu) = p_{\pm}(\mu)/\delta t$, the mean of the
102 time-averaged statistic satisfies

$$\mathbb{E}[\hat{g}] \approx \frac{\eta}{\delta t} \mu, \tag{S25}$$

103 and its variance obeys the scaling in Equation (S22). The validity of this local linear approximation is illustrated
104 in Supplementary Fig. SS3j, which highlights a near-linear response envelope around $\mu \approx 0$ under a locally
105 constant variance assumption.

Supplementary Note 4: Cramér-Rao lower bound of frame and event camera

Building on the statistical measurement model analyzed in Supplementary Note 3, we examine the fundamental estimation limits imposed by frame and event sensing under the same throughput constraints.

The Cramér–Rao lower bound (CRLB) provides a lower bound on the variance of unbiased estimators^{2–4}. Specifically, if $\hat{\Theta}$ is an unbiased estimator of a parameter Θ , then

$$\text{Var}(\hat{\Theta}) \geq \frac{1}{\mathcal{I}(\Theta)}, \quad (\text{S26})$$

where $\mathcal{I}(\Theta)$ denotes the Fisher information associated with the observation model.

Let \mathbf{b} denote observable data generated according to the likelihood model $p(\mathbf{b} | \Theta)$. The Fisher information is defined as

$$\begin{aligned} \mathcal{I}(\Theta) &= \mathbb{E} \left[-\frac{\partial^2}{\partial \Theta^2} \log p(\mathbf{b} | \Theta) \right] \\ &= \mathbb{E} \left[\left(\frac{\partial}{\partial \Theta} \log p(\mathbf{b} | \Theta) \right) \left(\frac{\partial}{\partial \Theta} \log p(\mathbf{b} | \Theta) \right)^\top \right] \\ &= \mathbb{E}[J(\Theta)J(\Theta)^\top], \end{aligned} \quad (\text{S27})$$

where

$$J(\Theta) = \frac{\partial}{\partial \Theta} \log p(\mathbf{b} | \Theta) \quad (\text{S28})$$

is the score function. The expectation is taken with respect to the distribution of the observable data \mathbf{b} .

When multiple independent measurements of the parameter Θ are available, the additive structure of Fisher information becomes explicit. Let $\{\mathbf{b}_t\}_{t=1}^T$ denote T independent observations drawn from $p(\mathbf{b}_t | \Theta)$. Under the iid assumption, the joint likelihood factorizes as

$$p(\{\mathbf{b}_t\} | \Theta) = \prod_{t=1}^T p(\mathbf{b}_t | \Theta). \quad (\text{S29})$$

The joint score function is therefore

$$\begin{aligned} J(\Theta) &= \frac{\partial}{\partial \Theta} \log p(\{\mathbf{b}_t\}_{t=1}^T | \Theta) \\ &= \sum_{t=1}^T \frac{\partial}{\partial \Theta} \log p(\mathbf{b}_t | \Theta) = \sum_{t=1}^T J_t(\Theta), \end{aligned} \quad (\text{S30})$$

where $J_t(\Theta)$ denotes the score contribution from the t -th measurement. Taking expectation yields the additive Fisher information identity

$$\mathcal{I}(\Theta) = \sum_{t=1}^T \mathcal{I}_t(\Theta). \quad (\text{S31})$$

Accordingly,

$$\text{Var}(\hat{\Theta}) \geq \frac{1}{\sum_{t=1}^T \mathcal{I}_t(\Theta)}, \quad (\text{S32})$$

indicating that repeated independent measurements reduce estimator variance through information accumulation.

124 **S4.1 CRLB of frame measurements**

125 In the frame modality, the observable data consist of photon counts $\{\mathbf{N}_t\}_{t=1}^T$ acquired independently at each
 126 exposure. The parameter of interest is the underlying scene intensity I , observed through the photon count $\lambda = I \delta t$
 127 accumulated over an exposure duration δt . For completeness, we summarize standard CRLB expressions for
 128 frame-based photon-counting measurements, which serve as a reference scaling under the throughput constraint.

129 **Poisson noise.** Under photon-limited conditions,

$$\mathbf{N}_t \stackrel{\text{iid}}{\sim} \text{Poisson}(\lambda). \quad (\text{S33})$$

130 The score function with respect to I is

$$J_t(I) = \frac{\partial}{\partial I} \log p(\mathbf{N}_t | I) = \delta t \left(\frac{\mathbf{N}_t}{\lambda} - 1 \right). \quad (\text{S34})$$

131 Using $\mathbb{E}[\mathbf{N}_t] = \lambda$ and $\text{Var}(\mathbf{N}_t) = \lambda$, the per-frame Fisher information becomes

$$\mathcal{I}_t(I) = \frac{\delta t}{I}. \quad (\text{S35})$$

132 By additivity,

$$\mathcal{I}(I) = \frac{T \delta t}{I}, \quad (\text{S36})$$

133 leading to the CRLB

$$\text{Var}(\hat{I}) \geq \frac{I}{T \delta t}. \quad (\text{S37})$$

134 **Mixed photon and readout noise.** When readout noise contributes additively, the photon-counting process
 135 can be approximated by

$$\mathbf{N}_t \stackrel{\text{iid}}{\sim} \mathcal{N}(\lambda, \lambda + \sigma_r^2). \quad (\text{S38})$$

136 Treating the variance as locally constant with respect to I , the per-frame Fisher information is approximated as

$$\mathcal{I}_t(I) \approx \frac{(\delta t)^2}{\lambda + \sigma_r^2}. \quad (\text{S39})$$

137 Summing over T frames yields

$$\mathcal{I}(I) \approx \frac{T(\delta t)^2}{I \delta t + \sigma_r^2}, \quad (\text{S40})$$

138 and the corresponding CRLB

$$\text{Var}(\hat{I}) \gtrsim \frac{I \delta t + \sigma_r^2}{T(\delta t)^2}. \quad (\text{S41})$$

139 These expressions provide a reference precision scaling for frame-based measurements under photon and readout
 140 noise.

141 **S4.2 CRLB of event measurements**

142 In the event modality, the observable data are accumulated bipolar event statistics $\{\hat{g}_t\}_{t=1}^T$ obtained by temporally
 143 aggregating motion-aligned events. The parameter of interest remains the scene intensity I , while the event
 144 statistics encode the motion-aligned log-intensity gradient

$$\mu \triangleq \mathbf{A}_v \log I.$$

145 **Gaussian approximation.** As shown in Supplementary Note 3, accumulated bipolar events are well
 146 approximated by a Gaussian distribution,

$$\hat{g}_t \sim \mathcal{N}(m(\mu), \sigma_g^2(\mu)), \quad \sigma_g^2(\mu) = \frac{p_+(\mu) + p_-(\mu)}{N}. \quad (\text{S42})$$

147 Treating the variance as locally constant with respect to μ , the score function becomes

$$J(\mu) \approx \frac{\hat{g} - m(\mu)}{\sigma_g^2(\mu)} m'(\mu). \quad (\text{S43})$$

148 The Fisher information with respect to μ is therefore

$$\mathcal{I}_\mu(\mu) \approx \frac{N (m'(\mu))^2}{p_+(\mu) + p_-(\mu)}. \quad (\text{S44})$$

149 In the near-linear regime where $m'(\mu) \approx \eta$, this yields the envelope

$$\mathcal{I}_\mu(\mu) \leq \frac{N \eta^2}{\alpha_0}, \quad \alpha_0 \triangleq p_+(0) + p_-(0), \quad (\text{S45})$$

150 corresponding to the maximum Fisher information attainable under the Gaussian approximation.

151 **Gradient envelope.** Since $\mu = \mathbf{A}_v \log I$, the Jacobian with respect to the intensity field is

$$\frac{\partial \mu}{\partial I} = \mathbf{A}_v \text{diag}\left(\frac{1}{I}\right). \quad (\text{S46})$$

152 The Fisher information for estimating I from event statistics therefore takes the quadratic form

$$\mathcal{I}(I) = \left(\frac{\partial \mu}{\partial I}\right)^\top \mathcal{I}_\mu(\mu) \left(\frac{\partial \mu}{\partial I}\right), \quad (\text{S47})$$

153 which explicitly shows that \mathbf{A}_v acts as a motion-aligned gradient envelope governing information transfer from I
 154 to event measurements. The corresponding event-based CRLB follows as $\text{Var}(\hat{I}) \gtrsim 1/\mathcal{I}(I)$ under the adopted
 155 approximation.

Supplementary Note 5: Working with conventional computational imaging solvers

Conventional computational imaging solvers typically describe the measurement process through a unified forward operator that maps a latent scene representation to observable measurements. This note explains how the NeuroSR formulation can be instantiated within existing computational imaging solvers. Our goal is not to redefine task-specific optical forward models, which are often well established in frame-based or other synchronous imaging modalities, but to clarify how the asynchronous event sensing likelihood can be *plugged into* conventional Maximum a posteriori (MAP)-based solvers in a compatible and modular manner. Throughout this note, the forward image formation model is assumed to be given, as specified in the main text. Here we focus on how this model interfaces with event-based constraints at the optimization level.

S5.1 Maximum a posteriori formulation

We consider a general MAP formulation in which the unknown quantity represents a latent scene variable, such as an intensity image or a task-dependent object/field. Let $I(\mathbf{x})$ denote a latent sharp intensity at a reference time (when applicable), and let θ denote additional parameters to be jointly optimized. Given measurements \mathbf{b} , MAP estimation takes the form

$$\begin{aligned}
 I^*(\mathbf{x}), \theta^* &= \arg \max_{I, \theta} p(I, \theta | \mathbf{b}) = \frac{p(\mathbf{b} | I, \theta) p(I) p(\theta)}{p(\mathbf{b})}, \\
 &\propto \arg \min_{I, \theta} \underbrace{-\log p(\mathbf{b} | I, \theta)}_{\text{Log-Likelihood}} + \underbrace{(-\log p(I) - \log p(\theta))}_{\text{Prior}} + \underbrace{\log p(\mathbf{b})}_{\text{Constant}},
 \end{aligned} \tag{S48}$$

where the first term encodes the measurement likelihood and the remaining terms impose priors or regularization. Different sensing modalities correspond to different likelihood terms, while sharing the same latent variables.

S5.2 Event-based MAP estimation

A characteristic property of event measurements is that motion and intensity can be efficiently decoupled at the likelihood level. We therefore first consider a motion-resolved ($\theta = \mathbf{v}$) MAP formulation based on the event channel alone. Assuming independent priors on I and \mathbf{v} , MAP estimation from event measurements \mathbf{b}_E yields

$$\begin{aligned}
 I^*(\mathbf{x}), \mathbf{v}^* &= \arg \max_{I, \mathbf{v}} p(I, \mathbf{v} | \mathbf{b}_E) \\
 &\propto \arg \min_{I, \mathbf{v}} -\log p(\mathbf{b}_E | I, \mathbf{v}) - \log p(I) - \log p(\mathbf{v}).
 \end{aligned} \tag{S49}$$

This formulation admits a natural two-stage instantiation, which is commonly adopted in event-based inference and is fully compatible with the NeuroSR framework.

1. **Motion retrieval from events.** The motion field \mathbf{v} can be estimated independently by leveraging the spatio-temporal contrast encoded in the event measurement \mathbf{b}_E . When the latent intensity I is unknown, \mathbf{b}_E alone serves as a motion-sensitive proxy constructed from warped events. In this setting, Equation (S49) can be interpreted probabilistically as selecting the flow that maximizes a decoupled likelihood term, with the intensity-dependent component replaced by a surrogate objective.

184 In practice, it is well established that correctly motion-compensated events exhibit strong spatial con-
 185 centration⁵, while incorrect motion hypotheses lead to temporally smeared and spatially diffuse event
 186 distributions. Using contrast maximization as an efficient surrogate, motion estimation reduces to

$$\mathbf{v}^* = \arg \min_{\mathbf{v}} \underbrace{-\text{Var}(\mathbf{b}_E(\mathbf{v}))}_{\text{Surrogate fidelity}} + \underbrace{\beta_{\mathbf{v}} \mathcal{R}(\mathbf{v})}_{\text{Flow regularizer}}, \quad (\text{S50})$$

187 where $\mathcal{R}(\mathbf{v})$ encodes spatial or temporal regularity of the flow field.

188 Importantly, the objective in Equation (S50) is independent of I and enables an efficient event-only pre-
 189 optimization of motion. This step provides a reliable initialization and constraint for subsequent joint
 190 intensity–motion refinement.

191 **2. Intensity retrieval.** Given the estimated motion \mathbf{v}^* , the event-channel likelihood simplifies considerably.
 192 As analyzed in Supplementary Note 3, temporal accumulation of motion-aligned events yields an approxi-
 193 mately Gaussian statistic whose mean is proportional to $\mathbf{A}_{\mathbf{v}^*} \log I$. Under this Gaussian approximation, the
 194 event likelihood can be written as

$$p(\mathbf{b}_E | I, \mathbf{v}^*) \propto \exp\left(-\mathcal{D}\{\mathbf{b}_E(\mathbf{v}^*), \mathbf{A}_{\mathbf{v}^*} \log I\}\right),$$

195 leading to the conditional MAP problem

$$I^*(\mathbf{x}) \propto \arg \min_I \underbrace{\mathcal{D}\{\mathbf{b}_E(\mathbf{v}^*), \mathbf{A}_{\mathbf{v}^*} \log I\}}_{\text{Event data fidelity}} + \underbrace{\beta_I \mathcal{R}(I)}_{\text{Image regularizer}}, \quad (\text{S51})$$

196 where the first term is also known as the data fidelity term, since it should provide a large proportion of the
 197 back-propagated gradient flow during optimization/training.

198 In principle, the data-fidelity term in the MAP formulation should be chosen to match the underlying event
 199 generation process. Assuming that each motion-compensated event e'_k is conditionally independent given
 200 the latent intensity $I(\mathbf{x})$ and the estimated flow \mathbf{v}^* , the likelihood of the aligned event cluster $\mathcal{E}' = \{e'_k\}_{k=1}^{N_e}$
 201 factorizes as $\prod_k p(e'_k | I)$. Accordingly, the time-averaged event measurement $\mathbf{b}_E = \sum_{k=1}^{N_e} e'_k / \Delta t$ admits a
 202 Gaussian approximation under temporal accumulation, *i.e.*, $p(\mathbf{b}_E | I) \sim \mathcal{N}(\boldsymbol{\mu}(e'_k), \text{Var}(e'_k) / N_e)$, where the
 203 mean response satisfies $\boldsymbol{\mu} \propto \mathbf{A}_{\mathbf{v}^*} \log I$, as derived in Supplementary Note 3. This Gaussian approximation
 204 provides a statistically consistent and solver-compatible likelihood for event measurements, despite the
 205 underlying event generation being nonlinear and asymmetric. An illustrative joint intensity–motion
 206 reconstruction example is shown in Supplementary Fig. Figure S4.

207 S5.3 Joint event-frame MAP and practical instantiation

208 NeuroSR combines the complementary constraints imposed by frame and event measurements by jointly
 209 optimizing their likelihoods under a shared latent intensity and motion field. Conditioning on the pre-estimated
 210 motion \mathbf{v}^* from Equation (S50), the resulting joint objective reduces to the event–frame MAP formulation
 211 presented in the main text, in which frame and event channels impose complementary likelihood terms on
 212 the same latent encoded intensity. This structure is directly compatible with existing frame-based solvers and

213 enables a plug-and-play integration of the event likelihood without modifying the underlying solver architecture.
 214 In practice, the motion field $\mathbf{v}(t)$ is parameterized as a piecewise-linear function and estimated from events
 215 using contrast maximization. Motion-compensated event aggregation specifies how the event measurement
 216 $\mathbf{b}_E(\mathbf{x}; \mathbf{v})$ is constructed in practice, with subpixel support achieved through pixel reassignment in IWE rendering
 217 (see Figure S5 and Supplementary Fig. Figure S4). These implementation details determine how the event
 218 likelihood enters the joint MAP formulation, while leaving the optimization structure unchanged.

219 We next summarize representative instantiations of this formulation across three computational imaging tasks:
 220 (i) large-FOV industrial inspection, (ii) biological microscopy screening, and (iii) snapshot pixel-SR lensless
 221 holography. The table below lists typical frame-domain operators, task-dependent optical forward models, and
 222 corresponding regularization choices used in these settings. These examples illustrate how the same joint MAP
 223 structure accommodates different imaging configurations.

Table S1. Examples of forward models and corresponding regularizers in typical application scenes. Here, \downarrow denotes the down-sampling operation, and \int denotes the temporal accumulation operation (long-exposure).

Setup	Sensor Model	f_{Θ}	\mathcal{R}
Industrial Inspection	\downarrow	Unity / Magnified Relay	2D Image Regularizer (TV)
Microscopic Screening	$\downarrow + \int$	Objective with Occlusion	2D Image Regularizer (TV, L1)
Pixel-SR Holography	$\downarrow + \int$	Free-space Propagation	Complex Field Regularizer (TV, L1)
More			

224 To accommodate different statistical assumptions or simplifications about the measurement noise, the
 225 likelihood terms in the joint MAP formulation can be instantiated with various negative log-likelihood losses, *i.e.*,
 226 the distance function $\mathcal{D}\{\cdot\}$. We summarize several commonly used choices in computational imaging below:

Table S2. Examples of negative log-likelihood losses under different statistical assumptions. Here, $\mathbf{a} = \{a_i\}$ denotes the prediction, and $\mathbf{b} = \{b_i\}$ the observation.

$\mathcal{L}(\mathbf{b}, \mathbf{a})$	Loss Name	Noise Model
$\ \mathbf{b} - \mathbf{a}\ _2^2$	L_2 loss	Gaussian distribution noise
$\ \mathbf{b} - \mathbf{a}\ _1$	L_1 loss	Laplace distribution
$\sum_i (a_i - b_i \log a_i)$	Poisson loss	Poisson distribution
More		

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235 with applications to motion, depth, and optical flow estimation. In *Proceedings of the IEEE conference on*
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S1 Supplementary Figures

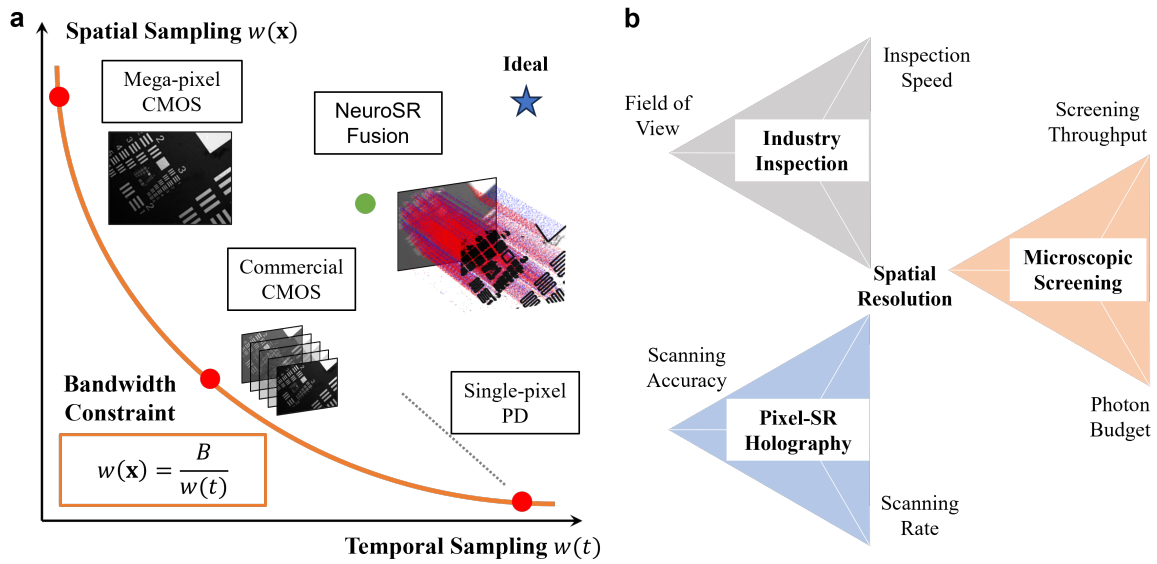


Fig. S1. Throughput-constrained sampling and application regimes. **a**, Illustration of the spatiotemporal throughput constraint in sensing, where spatial sampling density $w(\mathbf{x})$ and temporal sampling density $w(t)$ are coupled by a fixed budget B . Conventional frame-based imaging operates along the trade-off curve $w(\mathbf{x})w(t) = B$, linking spatial resolution to frame rate. NeuroSR leverages asynchronous, motion-induced sampling to redistribute measurements within this constraint. **b**, Projection of representative high-throughput imaging tasks into the resolution–throughput space, including industrial inspection, microscopic screening, and pixel super-resolution holography, highlighting their distinct operating regimes under shared throughput limitations.

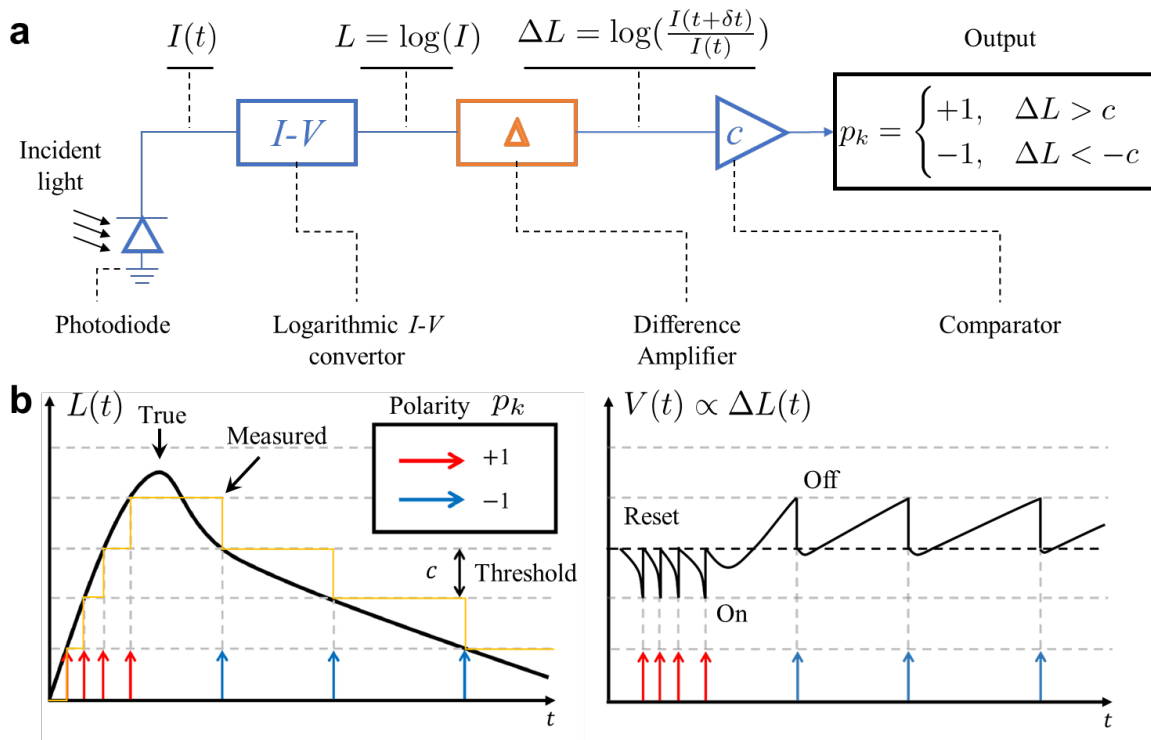


Fig. S2. Asynchronous triggering mechanism of an event camera. **a**, Circuit-level principle of event generation: a pixel reports an event when the logarithmic intensity change exceeds a fixed contrast threshold c . **b**, Upon threshold crossing, the pixel voltage is reset and a polarity-coded event is emitted, encoding the sign of the log-intensity increment.

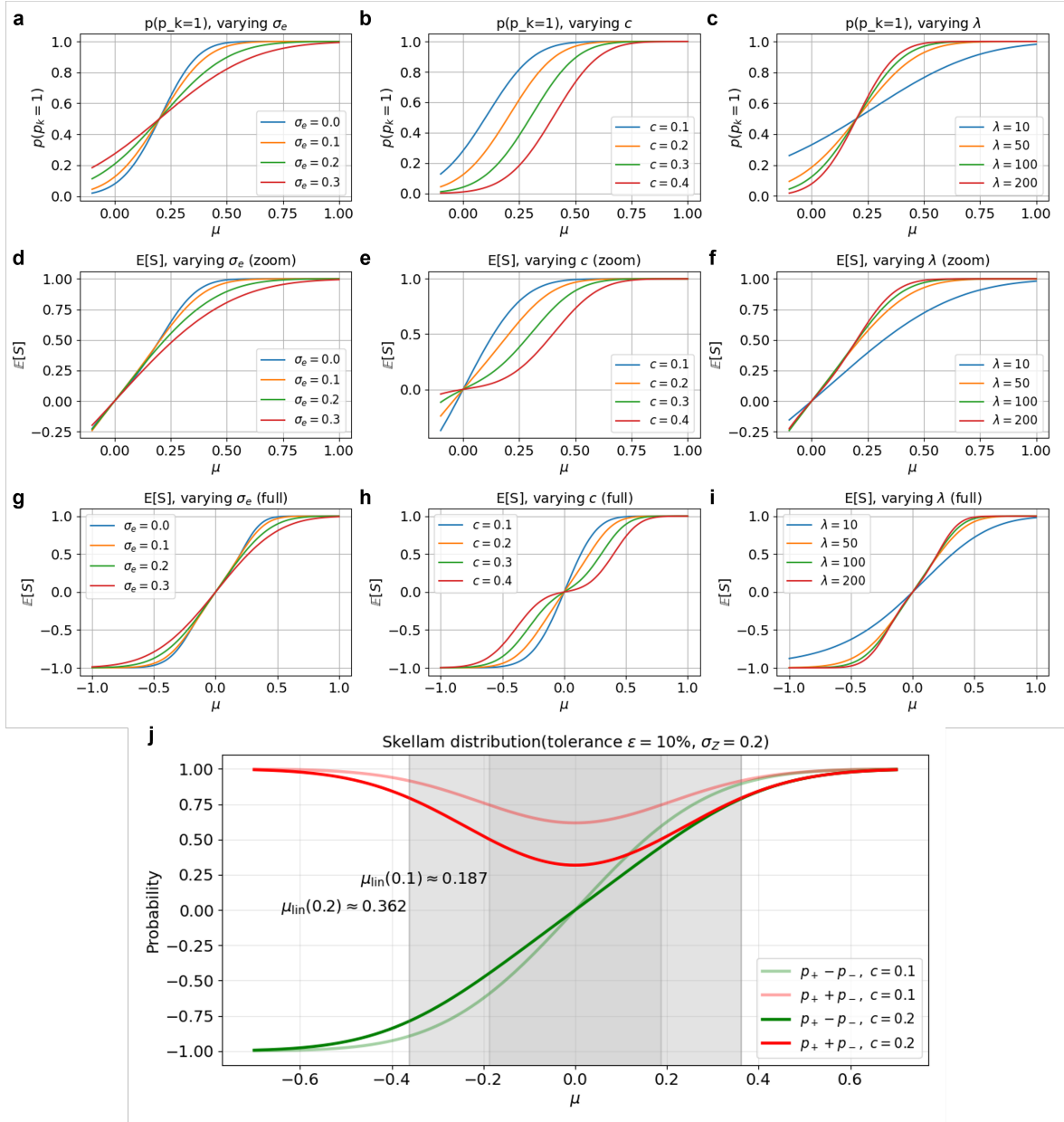


Fig. S3. Event triggering statistics and near-linear bipolar response. **a-c**, Single-polarity probability $p(p_k = 1)$ as a function of the motion-aligned increment μ under variations of readout noise σ_e , threshold c , and photon level λ (zoomed range). **d-f**, Expected bipolar polarity $\mathbb{E}[S]$ in the same zoomed range, showing reduced sensitivity near $\mu \approx 0$. **g-i**, Full-range $\mathbb{E}[S]$ curves, exhibiting stable sigmoidal transitions. **j**, Linearization envelope under a locally constant variance approximation (σ_z fixed), where the shaded interval indicates the near-linear regime $|\mu| \leq \mu_{\text{lin}}(c)$ defined by a relative linearization error tolerance.

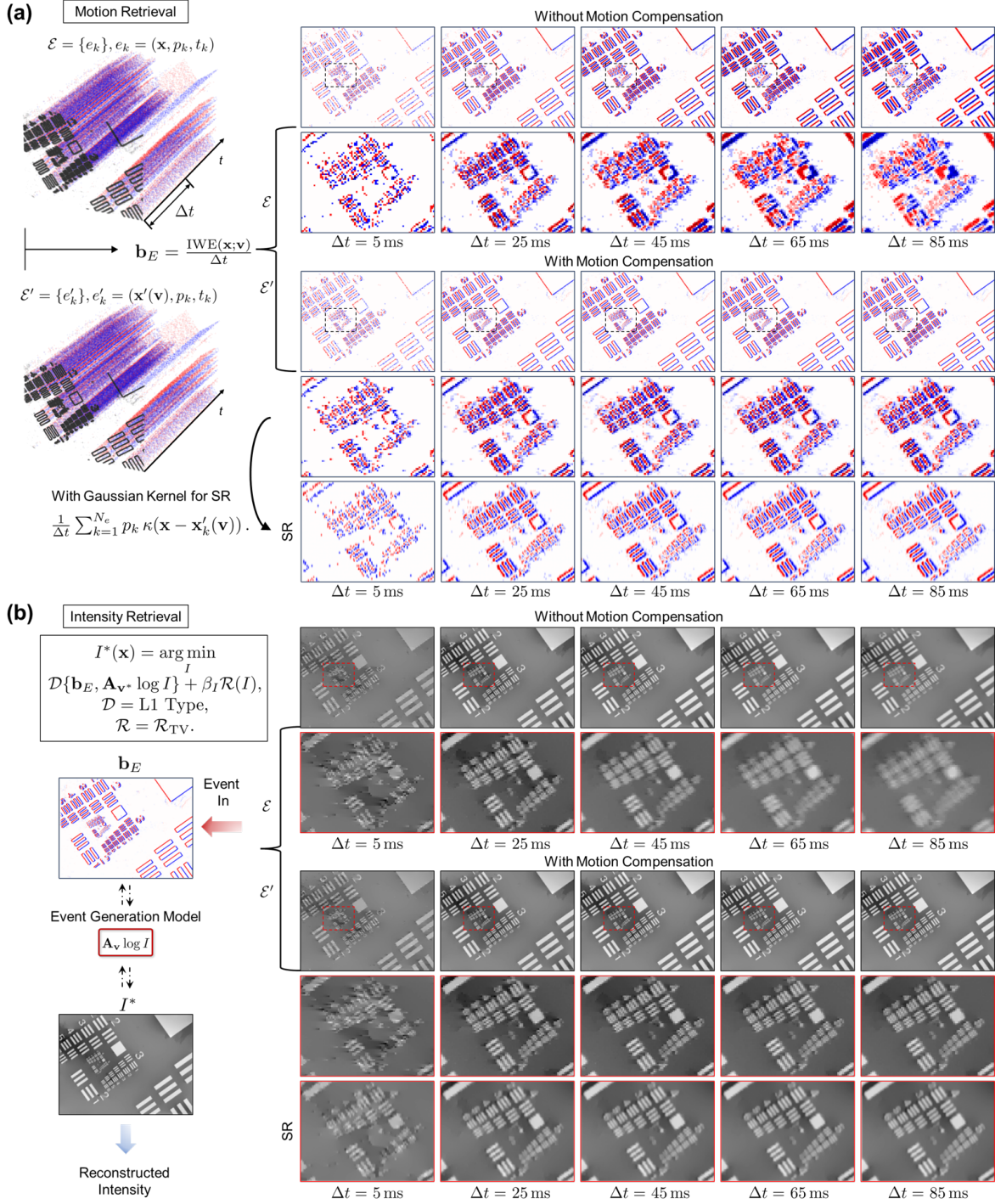


Fig. S4. Asynchronous feature encoding. **a**, Warping events along the estimated flow \mathbf{v}^* reassigns temporally separated event samples to consistent sub-pixel spatial locations at a common reference time, enabling coherent accumulation of sub-pixel information over an extended temporal window. **b**, Under correct motion compensation, increasing the accumulation duration Δt improves statistical stability. Apparent smearing arises only from motion mismatch or model violation. The observed reconstruction sharpening reflects the recovery of sub-pixel structure. Bottom row (Zoom-in).

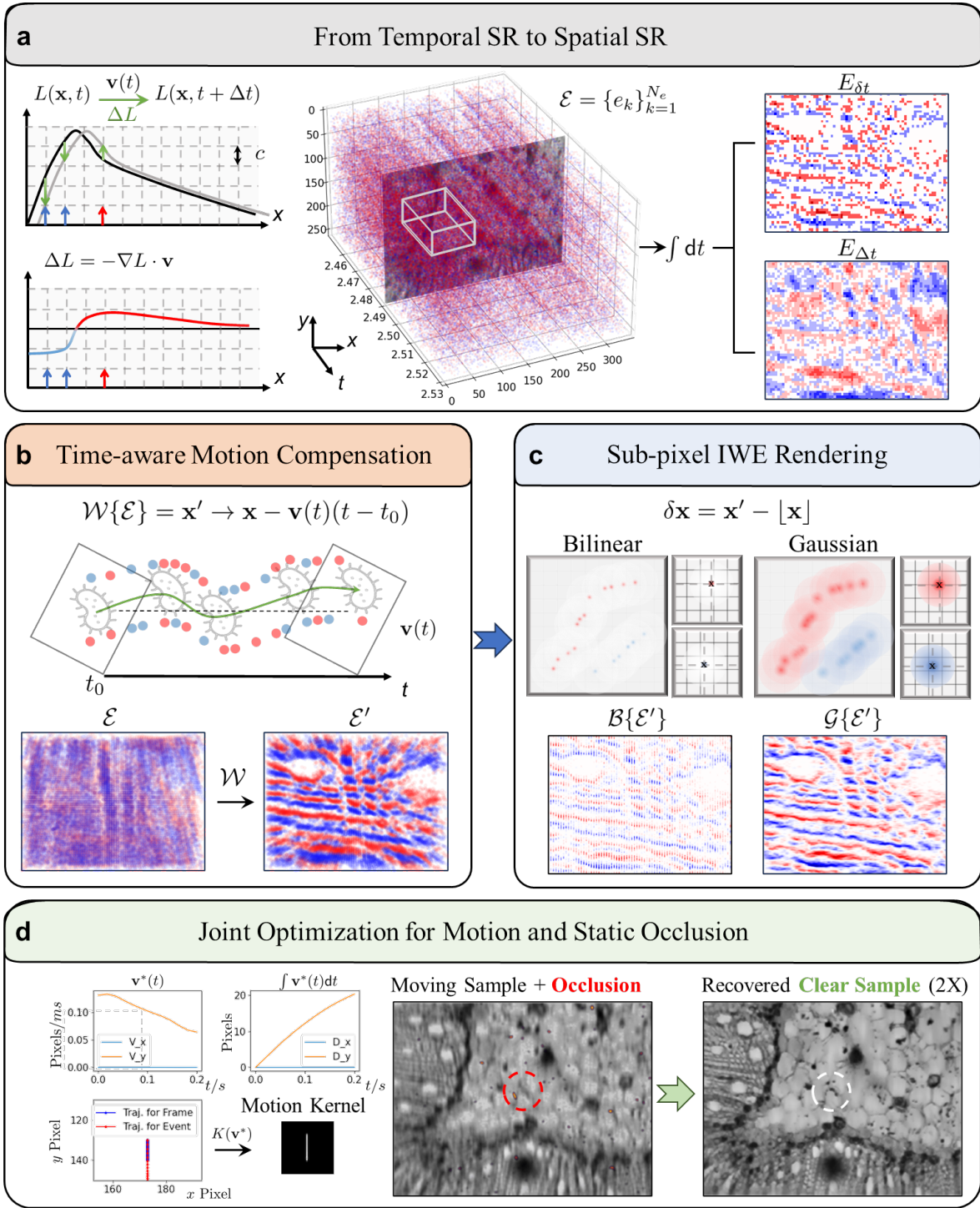


Fig. S5. Event-based pixel reassignment for sub-pixel-resolved IWE rendering. **a**, Conceptual illustration of how motion-induced temporal variations encoded by events are converted into spatially resolved information through integration. **b**, Motion-compensated warping of events to a common reference time using a time-varying flow model. **c**, Sub-pixel IWE construction, comparing discrete interpolation and continuous kernel-based rendering. **d**, Joint refinement of motion, occlusion and intensity enables deblurring and defect separation.