

# Supporting Information

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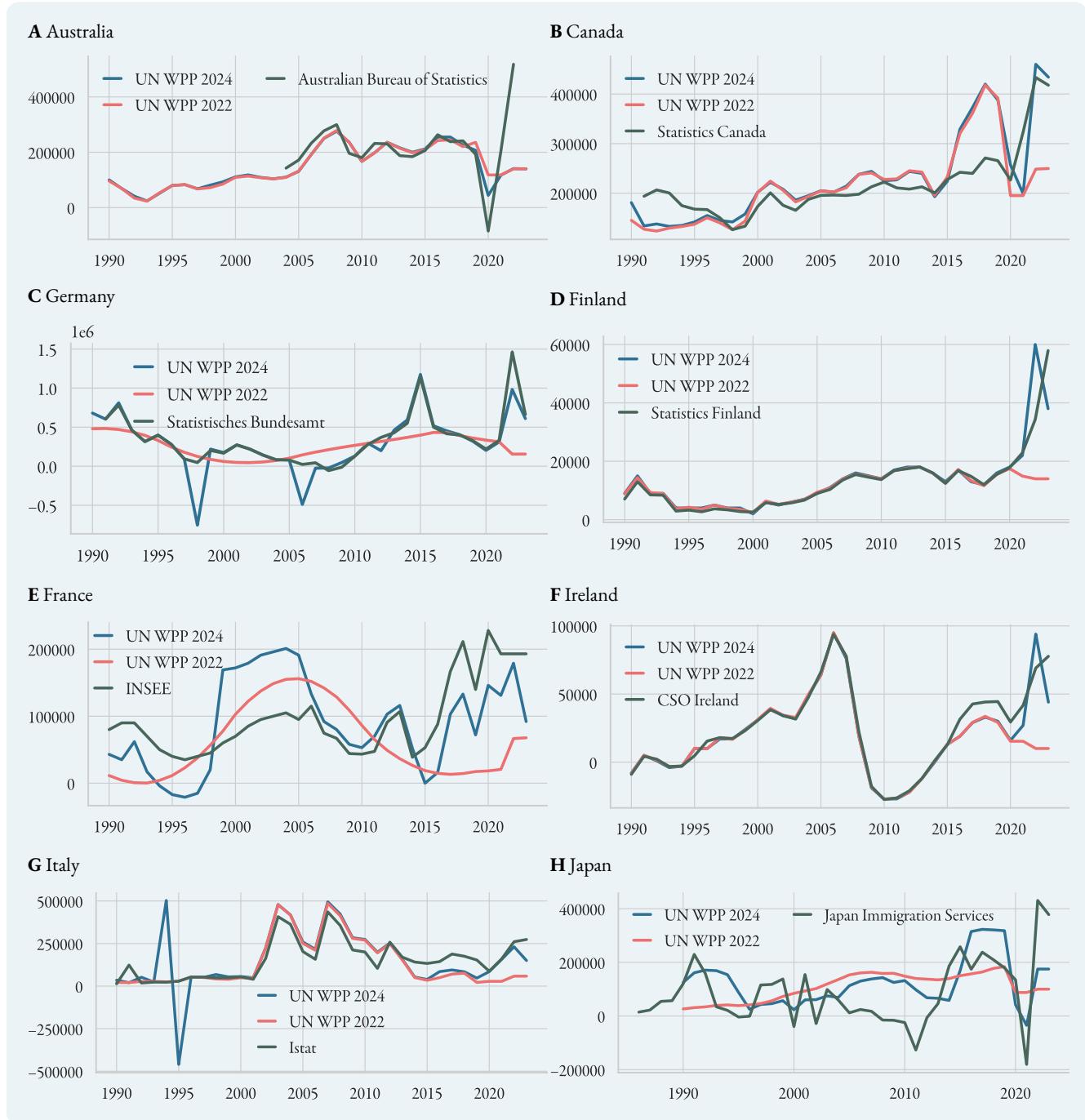
In the following we detail collection and assembly of the input and target data used to train the neural network, as well the validation process used to select the neural network architecture.

## Target data

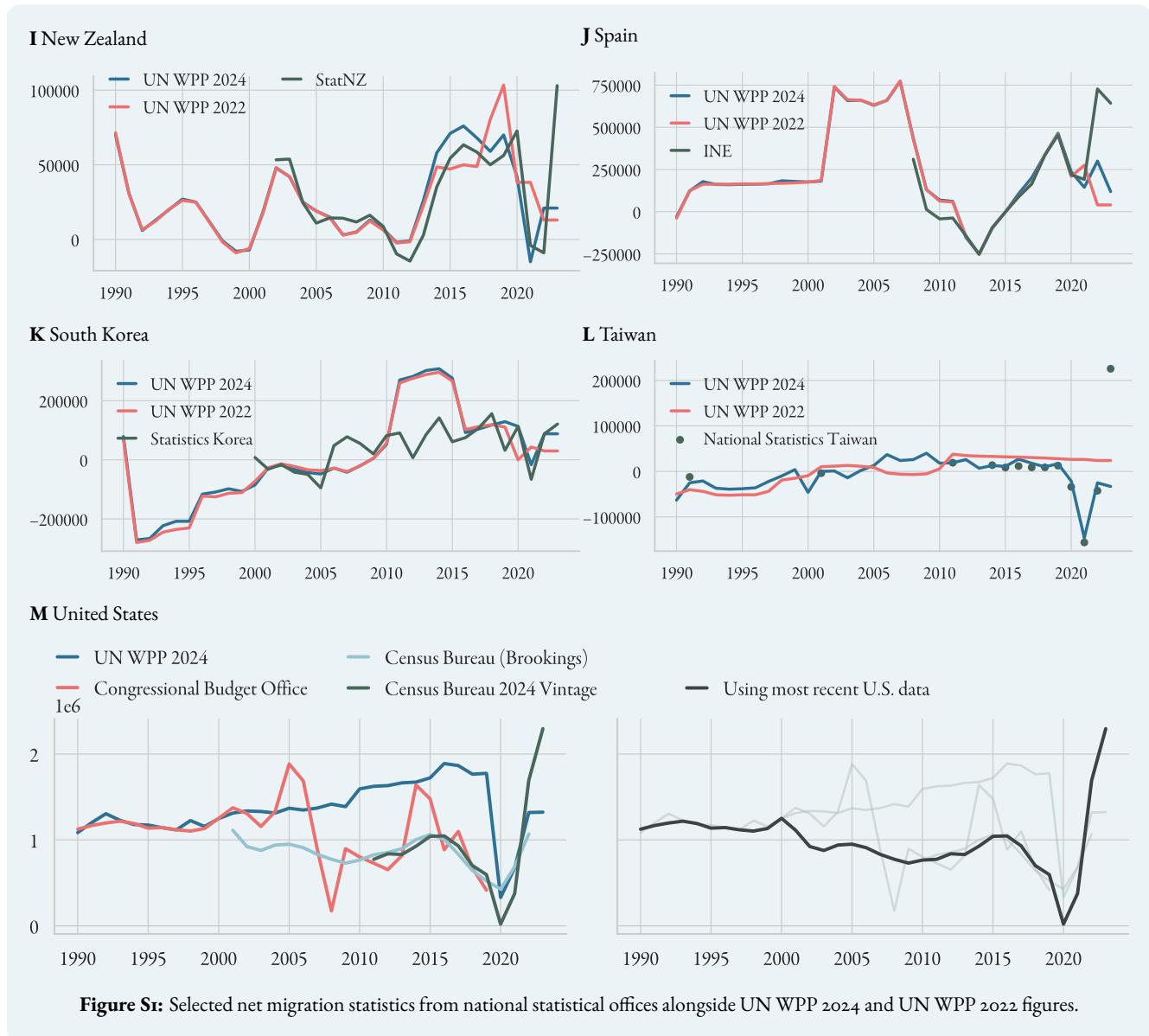
### Net migration data

To train the neural network, we use target net migration statistics for a small number of countries. This data are mostly taken from National Statistics Bureaus, as well as the UN 2024 World Population Prospects for a handful of additional countries and periods. In figures [S1](#) we show national figures alongside the UN WPP 2024 and WPP 2022 Revision data for some selected countries. UN WPP 2024 figures generally agree well with national estimates for periods pre-2010; larger discrepancies typically occur around 2020 due to the pandemic, the war in Ukraine, and a probable lack or lag in new demographic data reaching demographers at UN DESA.

- Australia: data from 2004 onwards provided by the Australian Bureau of Statistics [\[1\]](#). Pre-2004 we use UN WPP 2024 data.
- Austria: data from 2002 onwards provided by Statistik Austria [\[2\]](#). Pre-2002 we use UN WPP 2024 data.
- Belgium: data from 2000 onwards provided by StatBEL [\[3\]](#). Pre-2000 we use UN WPP 2024 data.
- Bulgaria: data from 2010 onwards provided by the National Statistical Office of Bulgaria [\[4\]](#).
- Canada: data provided by Statistics Canada [\[5\]](#). Since figures are given for years ending in the summer, we use the midpoint average as an approximation for calendar-year net migration figures.
- Czech Republic: data provided by the Czech Statistical Office [\[6\]](#).
- Denmark: data provided by Statistics Denmark [\[7\]](#).
- Estonia: data from 2000 onwards provided by Statistics Estonia [\[8\]](#); since the figures agree with the UN WPP 2024 data, pre-2000 we use UN WPP 2024 data.
- Finland: data provided by Statistics Finland [\[9\]](#).
- France: data are provided by INSEE (Institut National de la statistique et des études économiques) [\[10\]](#).
- Germany: data are provided by the Statistisches Bundesamt [\[11\]](#).
- Iceland: data provided by Statistics Iceland [\[12\]](#).
- Ireland: data provided by the Central Statistics Office [\[13\]](#); these are counted from April to April and thus need to be shifted back and interpolated accordingly.
- Italy: data from Istat [\[14, 15\]](#).
- Japan: data are provided by the Japan Immigration Services Agency [\[16\]](#).
- Latvia: data are provided by the Statistics Agency of Latvia [\[17\]](#).
- Lithuania: data are provided by Statistics Lithuania [\[18\]](#).

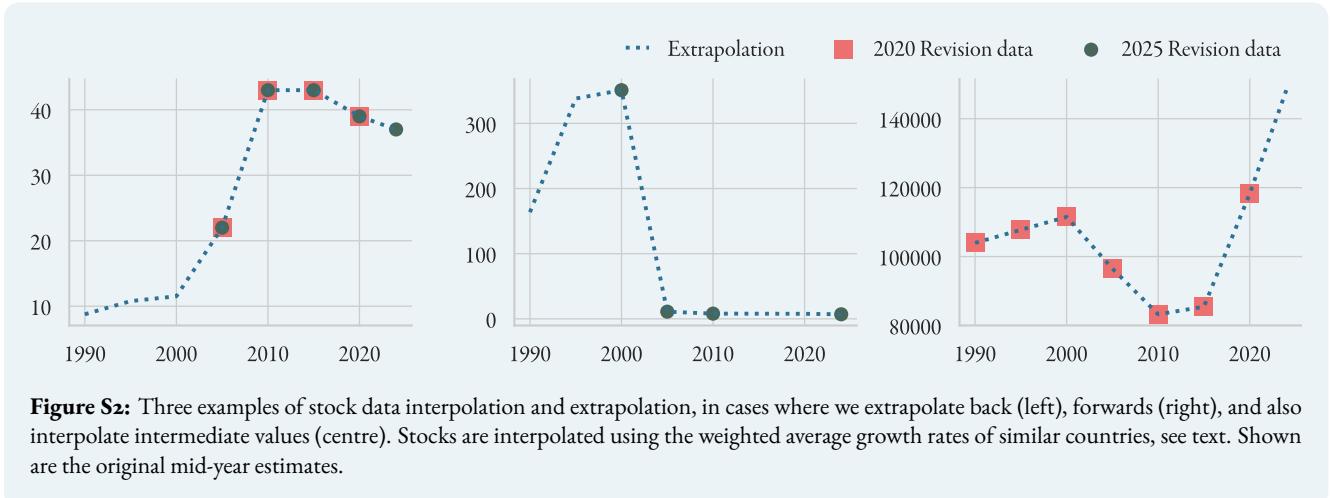


- Netherlands: data from 2003 onwards is provided by Statistics Netherlands [19] agree with the UN WPP 2024 figures, and these are thus used for the remaining years.
- New Zealand: data from 2002 onwards is provided by Statistics New Zealand [20]. Pre-2002 we use UN WPP data.
- Norway: data are provided by Statistics Norway [21].
- Portugal: data from 2002 onwards provided by Statistics Portugal [22]. Pre-2002 we use UN WPP 2024 data.



**Figure S1:** Selected net migration statistics from national statistical offices alongside UN WPP 2024 and UN WPP 2022 figures.

- Slovenia: data are provided by the Statistical Office of Slovenia [23].
- South Korea: Statistics Korea [24] publishes net migration figures from 2000 onwards. Pre-2000 values are not used.
- Spain: data from 2008 onwards is provided by INE (Instituto Nacional de Estadística) [25]. Pre-2008 we use UN WPP 2024 data.
- Sweden: data provided by Statistics Sweden [26].
- Switzerland: data are provided by the Bundesamt für Statistik [27] from 1991; for 1990, we use the UN WPP 2024 figure.
- Taiwan: data for some years is provided by National Statistics Taiwan [28]. Missing years are filled using the UN WPP 2024 data.



- United Kingdom: data are provided by the Office for National Statistics [29].
- United States: figures from the US Census Bureau [30, 31], the Congressional Budget Office [32], and the UN WPPs offer inconsistent pictures (see fig. S1M). We collate the data by always using the most recent figures from US bureaus, shown on the right.

## Migrant stocks

**Interpolation and Extrapolation** The stocks are both a target value and an input to the neural network. We use the UN DESA 2025 Revision data where available, and fill gaps using values from the 2020 Revision. For edges  $(i, j, k)$  where stocks  $S_{ij}(t)$  or  $S_{ik}(t)$  are missing for *all* time points  $t$ , we set the flow  $T_{ijk}(t) = T_{ikj}(t) = 0$  for all  $t$  if  $i \neq j$  or  $i \neq k$ , and set the initial stock value to 0. This is problematic as it underestimates some fairly major edges (e.g. Argentina or Brazil to India), but there is no stock data available from which an inference could be made. However, there are many series  $S_{ij}(t)$  where data are partially observed (i.e. values are missing for some, but not all time points  $t$ ). For these series, we extrapolate missing stocks in the following way: let  $c_{jk} = \text{corr}(S_{ij}, S_{ik})$  and  $\rho_{jk} = \exp(-d_{jk})$ , where  $d_{jk}$  is the geodesic distance between country  $j$  and  $k$ .

We want to extrapolate missing stock values by comparing them to ‘similar’ countries with complete data. ‘Similarity’ here means that (1) the stock time series correlate strongly, and (2) the countries are geographically close. We thus build a weight measure for a country pair  $(j, k)$  using

$$w_{jk} = c_{jk} \times \frac{\rho_{jk}}{\sum_k \rho_{jk}}.$$

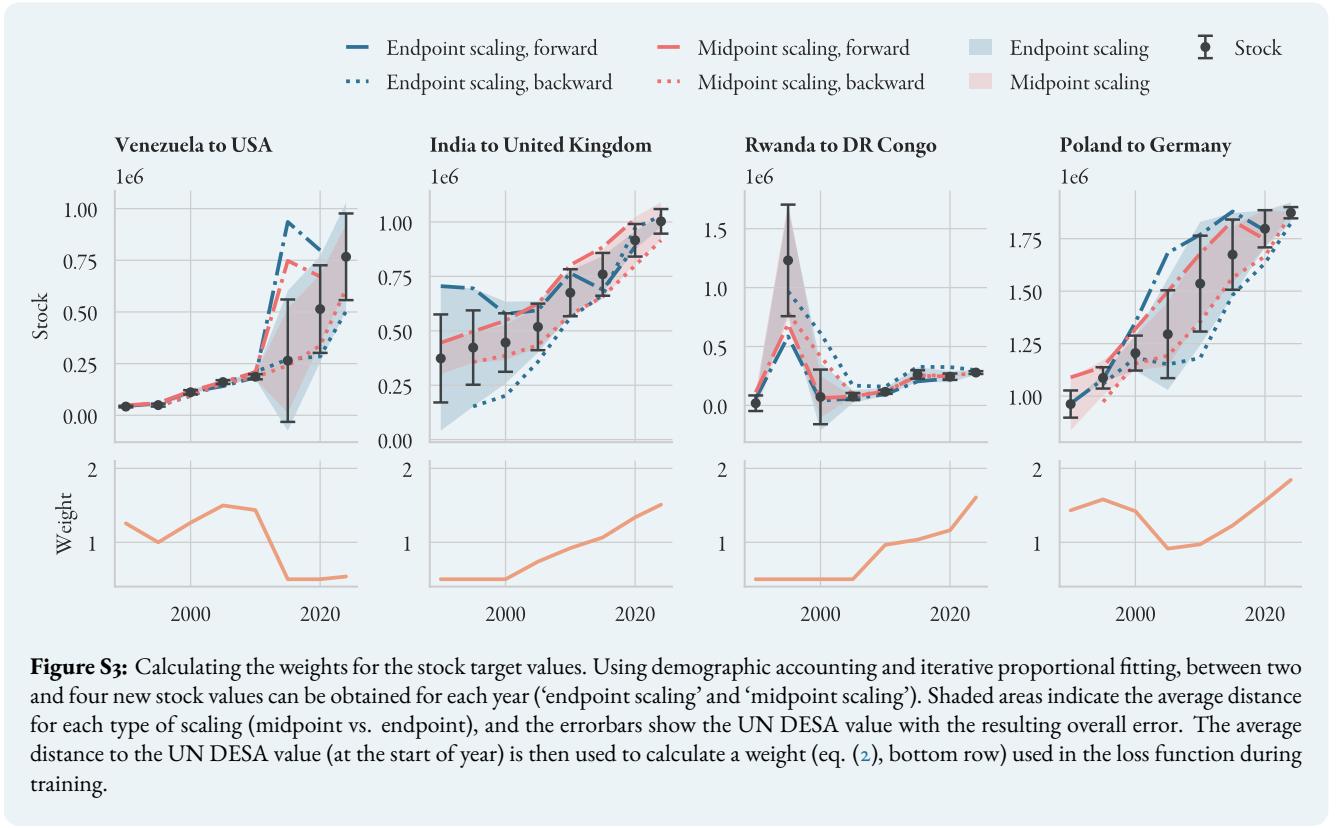
Then, given a matrix of stock growth values  $S_{ij}/S_{ij}(t-1) - 1$ , we define an average growth rate for the missing series as

$$\bar{g}_{ij} = \frac{\sum_k g_{ik} w_{jk}}{\sum_k w_{jk}}.$$

The missing stocks are then extrapolated and interpolated using this growth rate. This can be done even in the case of only a single datapoint (see fig. S2).

After interpolation, the native-born population  $S_{ii}$  is then estimated as

$$S_{ii}(t) = P_i(t) - \sum_{j \neq i} S_{ji},$$



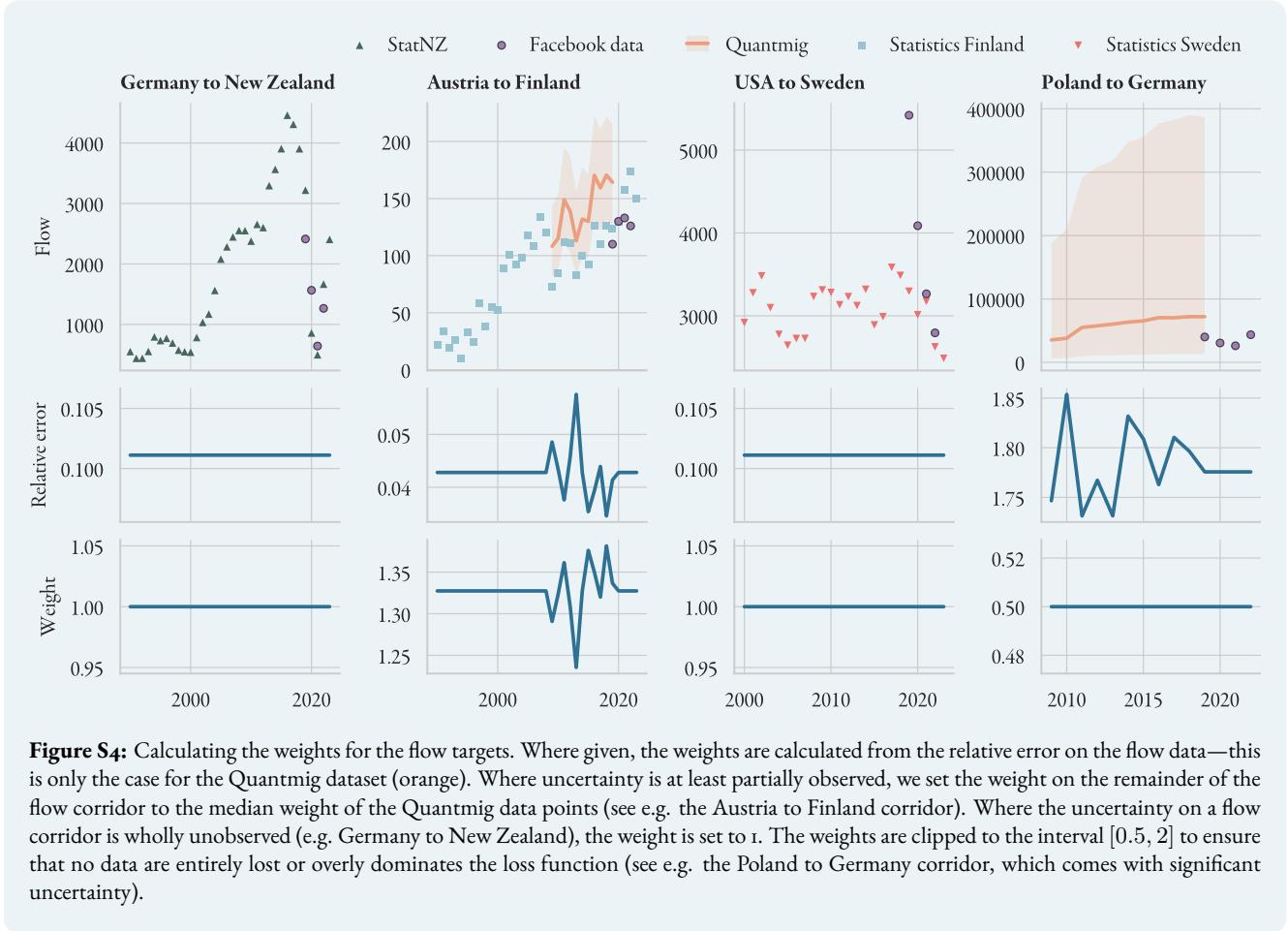
where  $P_i(t)$  is the total population (in July of each year), taken from the UN WPP dataset. The native-born population naturally makes up the upper tail of the migrant stock distribution (see fig. S16H).

$$\sum_i \frac{S_{ij}}{\sqrt{1 - \gamma_j}} - \frac{B_i}{2\sqrt{1 - \gamma_i}},$$

where the factor of 2 is a result of assuming that births are distributed evenly throughout the year, i.e. half the births occur in the first six months of each year. The column marginals represent the total population in January of each year, as given by the UN WPP dataset [33].

**Calculating the weights** We generate demographically closed stock matrices using iterative proportional fitting. Let  $S_1 = S(t_1)$ ,  $S_2 = S(t_2)$  be two successive stock tables, i.e.  $t_2 = t_1 + 5$ . In principle, adding the total number of births in the period  $[t_1, t_2]$  to  $S_1$  and the total deaths in  $[t_1, t_2]$  to  $S_2$  should result in stock matrices with the same marginals  $a = \sum_i S_{ij}$ ,  $b = \sum_j S_{ij}$ . In practice, this is not the case. We can thus calculate the marginals  $a_1, b_1, a_2, b_2$  for each stock matrix with births and deaths added, and scale each to the mid-point  $\frac{1}{2}(a_1 + a_2)$ ,  $\frac{1}{2}(b_1 + b_2)$  using IPF ('midpoint scaling'). After scaling, we subtract births and deaths again, obtaining new estimates of the stock table for each year. Doing this for each pair of stock tables, we obtain two estimates of the stock for each year except the years at the boundaries (1990 and 2024), for which we only have one; one value comes from comparison with the previous stock table ('backward') and one from comparison with the next table ('forward'). Additionally, we can also add births *and* subtract deaths *only* from  $S_1$ , and repeat the same procedure, obtaining two more estimates ('endpoint scaling'). In total, we thus obtain four additional estimates for each year except the boundary years, for which we have two (see fig. S3). The error on each stock value is then given by

$$\sigma_{ij} = \langle |S_{ij}^k - S_{ij}| \rangle_k, \quad (1)$$



where  $\langle \cdot \rangle_k$  is the average over the various estimates obtained from IPF. The weights are calculated based on the relative error  $\rho_{ij} = \sigma_{ij}/S_{ij}$ , where we first normalise  $\rho_{ij}$  to have mean 0 and variance 1; this ensures that the weights are balanced around 1. We also want to avoid excessive distortion of the loss function due to extremely large or extremely small weights, and hence truncate the weights to  $[0.5, 2]$ . The weights are thus given by

$$w_{ij} = \min \left( 2, \max \left( 0.5, \exp \left( -(\rho_{ij} - \langle \rho_{ij} \rangle) / \langle \rho_{ij}^2 \rangle^{1/2} \right) \right) \right), \quad (2)$$

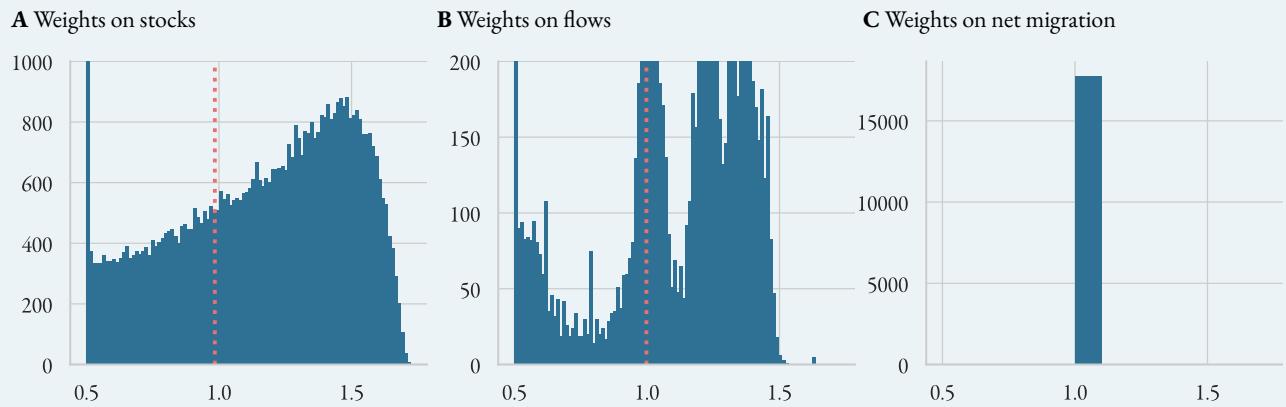
see fig. S5A. We estimate the error on a stock difference  $S_{ij}(t_2) - S_{ij}(t_1)$  by assuming the errors are independent, setting  $\sigma(\Delta S) = \sqrt{\sigma_1^2 + \sigma_2^2}$ , and calculating the weights as above.

## Flow data

The uncertainty on the flow targets is based on the uncertainty from the Quantmig estimates, which come with standard errors. As with the stocks, the weights are then given by

$$w_{ij} = \min \left( 2, \max \left( 0.5, \exp \left( -(\rho_{ij} - \langle \rho_{ij} \rangle) / \langle \rho_{ij}^2 \rangle^{1/2} \right) \right) \right), \quad (3)$$

with  $\rho_{ij}$  the relative error (standard error divided by central estimate). For all flow corridors along which we partially know the median errors, we set the relative error to the median relative error on that corridor, and calculate the weights as above.



**Figure S5:** Distribution of the weights on the stock, flow, and net migration target values. The distributions are clipped to  $[0.5, 2]$ , and each have mean 1 (red dotted lines). For the net migration, we currently set all weights to 1.

Finally, for flow corridors with entirely unobserved errors, we set the relative error to the mean of the Quantmig dataset and the weight to 1 (see fig. S4).

## Input data

### GDP per capita and GDP annual growth

In this section we detail the assembly of the economic covariates used to train the neural network: GDP per capita, in real 2015 USD; and GDP % annual growth, given by

$$\partial_t \text{GDP}(t) = 100 \times \left( \frac{\text{GDP}(t)}{\text{GDP}(t-1)} - 1 \right). \quad (4)$$

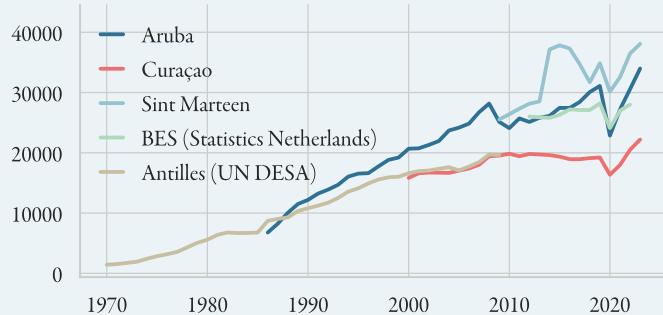
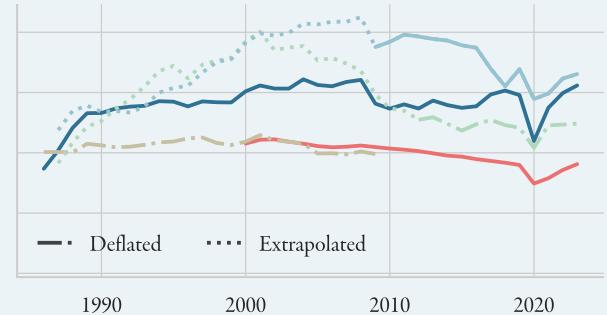
GDP data are sourced from the World Bank [34, 35] as well as UNCTAD [36]. However, these datasets are not complete for all countries and years required. Historical GDP growth rates can be calculated from the Maddison dataset [37] (giving GDP/capita in 2011 PPP) and the IMF dataset [38] (GDP/capita in 2021 PPP). Note that GDP growth can be calculated from any measure of real GDP (e.g. constant USD or constant PPP). GDP growth, when available, can be used to calculate GDP/capita for missing years. At times, only nominal GDP figures are available for a given year and country. Converting to real GDP requires the *deflator* for that country,

$$\text{GDP deflator}(t) = \frac{\text{nominal GDP}(t)}{\text{real GDP}(t)}. \quad (5)$$

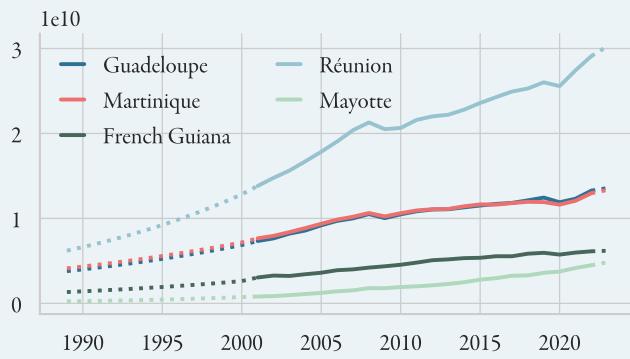
Where the deflator is unavailable, deflators from economically, geographically, or socially similar countries can be used to estimate real GDP. We also fill missing values using figures from national statistical bureaus and additional sources, as below.

GDP data are scaled using a Yeo-Johnson transform (see fig. S16A–B).

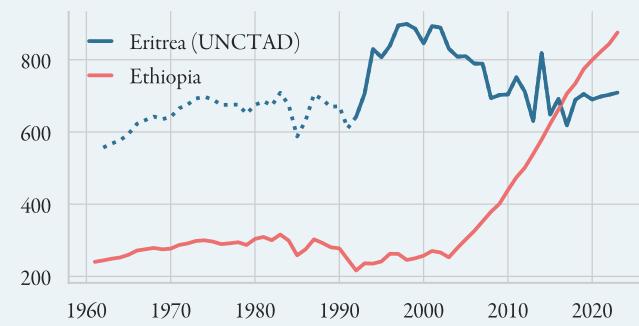
**Caribbean Netherlands** The Netherlands Antilles were a constituent country of the Kingdom of the Netherlands, consisting of the islands of Bonaire, St. Eustatius, and Saba (the BES islands), Aruba, Curaçao, and Sint Marteen. Aruba became an independent constituent country in 1986, and in 2010, Curaçao and Sint Marteen followed suit, at which point the Netherlands Antilles were dissolved and the remaining BES islands became a special municipality within the Kingdom.

**A** Nominal GDP per capita**B** Real GDP per capita, in const. 2015 USD

**Figure S6:** Caribbean Netherlands. **A** Available nominal GDP per capita figures for the constituent islands. **B** We fill missing years by deflating nominal GDP figures using the deflator for Aruba, and then extrapolate back using the growth rate from Aruba.



**Figure S7:** French overseas territories real GDP, in 2015 USD. The available data from Statista is extrapolated to the entire period 1989–2023 using the average growth rates for each country.



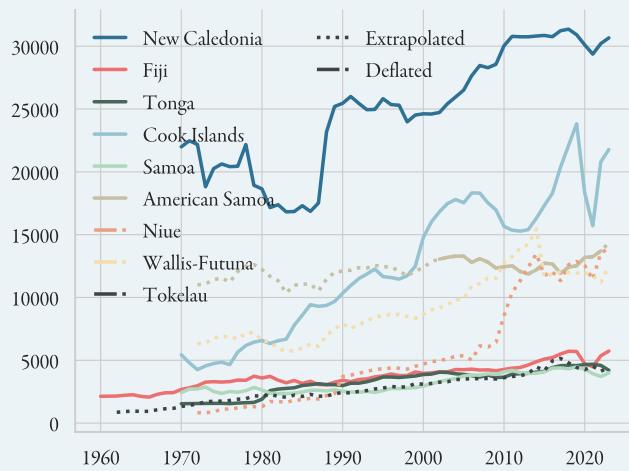
**Figure S8:** Eritrea. Since the country did not formerly achieve independence from Ethiopia until 1991, we extrapolate the available data from UNCTAD back to 1989 using the Ethiopian growth rate (dotted line).

The World Bank and UNCTAD provide real GDP figures for Aruba from 1986, for Curacao from 2000, and for Sint Maarten from 2009. Statistics Netherlands [39] provides nominal GDP for the BES islands from 2012–2022, while UN DESA provides nominal GDP for the entire Netherlands Antilles from 1970–2009 [40], which agrees with the World Bank data for Curaçao, the largest of the remaining Antilles islands after 1986 (see fig. S6A). Since the deflators are not available, we convert the UN DESA data to real GDP using the deflator for Aruba, and then extrapolate back to 1986 using the growth rate for Aruba (see fig. S6B).

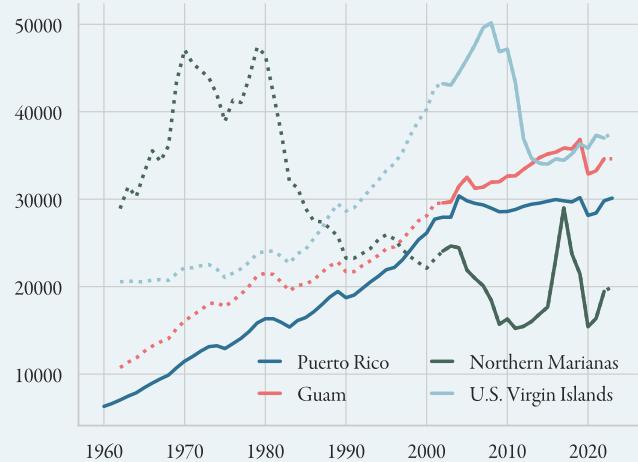
**French overseas territories** Statista [41] provides real GDP in constant 2014 Euros for the period 2000–2022. We convert to 2015 USD and extrapolate back to 1989 using the average growth rate for each territory, calculated over the period from 2000–2008. We extrapolate forward to 2023 using the average growth rate for each country in the period 2017–2022 (see fig. S7).

**Eritrea** Eritrea formally achieved independence from Ethiopia in 1991. UNCTAD provides data from 1992, and we extrapolate back to 1989 using the growth rate from Ethiopia (see fig. S8).

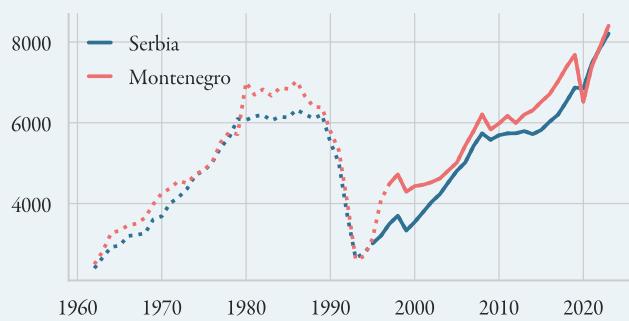
**Small Pacific Islands** The World Bank and UNCTAD datasets have missing data for American Samoa, Niue, Tokelau, Wallis and Futuna, Guam, and the Northern Mariana Islands. The Pacific Community provides nominal GDP figures



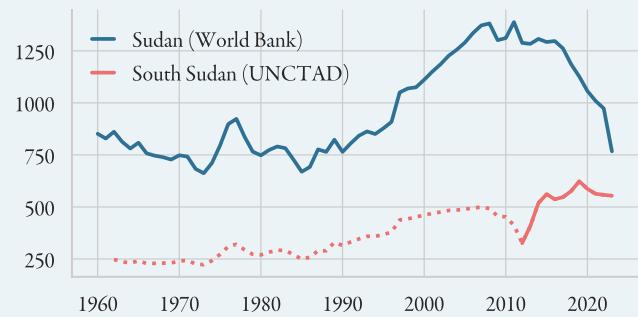
**Figure S9:** Small Pacific Islands GDP per capita, in const 2015 USD. Nominal GDP data from the Pacific Community data are deflated using the inflation rate as a proxy (dash-dotted lines). We then extrapolate back (dotted lines) using growth rates from neighbouring countries with similar GDP/capita.



**Figure S10:** US overseas territories GDP per capita, in const 2015 USD. Missing values for Guam, the Northern Marianas, and the US Virgin Islands are extrapolated back using the growth rate of Puerto Rico (dotted lines).



**Figure S11:** Serbia and Montenegro, GDP per capita in const 2015 USD. We extrapolate back to 1960 using the growth rates given by the Maddison project dataset (dotted lines).

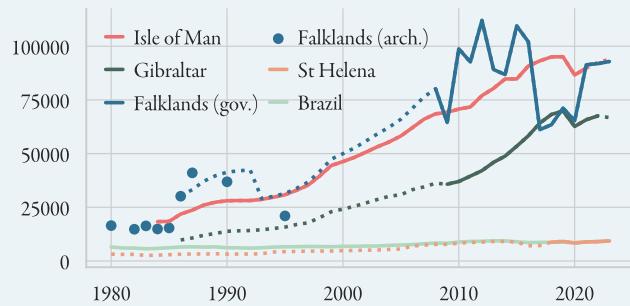


**Figure S12:** Sudan and South Sudan GDP per capita in const 2015 USD. We extrapolate the UNCTAD data for South Sudan back from 2011, the year of its independence, using the growth rate of Sudan (dotted line).

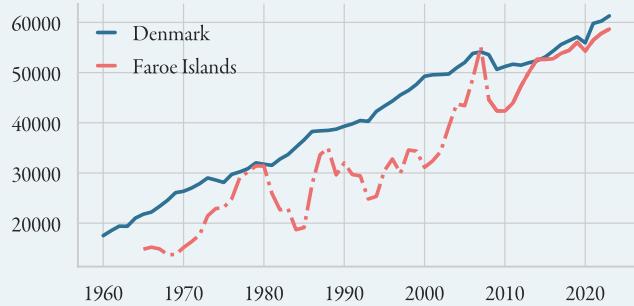
[42], as well as country-specific inflation rates [43], which we use as a proxy for the GDP deflator. We then extrapolate back using growth rates of neighbouring countries:

1. American Samoa, Niue, and Wallis-Futuna: average growth rate of New Caledonia, the Cook Islands, and Fiji,
2. Tokelau: growth rate of Fiji.

**US overseas territories** Guam and the Northern Marianas are US overseas territories that to a large part rely on US funding for their economies. We extrapolate both Guam and the Marianas using the growth rate from Puerto Rico. The US Virgin Islands we extrapolate back using the average growth rate of Puerto Rico, St. Kitts and Nevis, and Antigua and Barbuda.



**Figure S13:** British territories real GDP per capita in const 2015 USD. Missing data for Gibraltar and Falklands is extrapolated back from Government data using the Isle of Man growth rate (dotted lines). The Falklands extrapolation agrees broadly with archival data (blue dots). Government data for St. Helena is extrapolated back using the growth rate for Brazil, also shown for comparison purposes.



**Figure S14:** Faroe Islands real GDP per capita in const 2015 USD. Missing data for the islands is calculated by deflating the available nominal figures using the Danish deflator as a proxy (dash-dotted line).

**Serbia and Montenegro** The state of Serbia and Montenegro existed from 1992 until 2006, at which point the two constituent republics separated and became independent nations. Growth rates for these two countries are estimated in the Maddison project dataset, and are used to extrapolate GDP figures back to 1960 (see fig. S11).

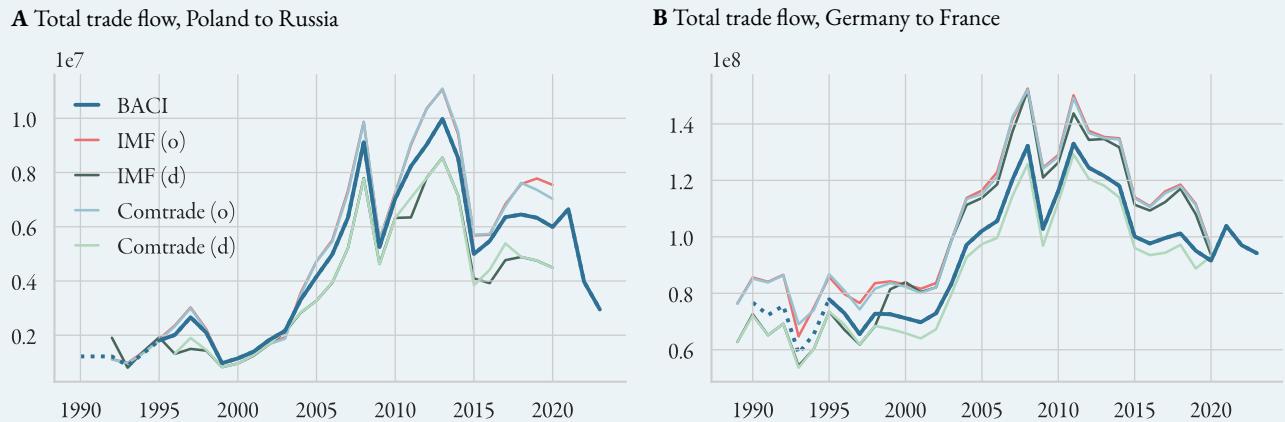
**South Sudan** South Sudan ceded from Sudan in 2011. UNCTAD gives real GDP figures from 2011 onwards, and we extrapolate back in time using the growth figures from Sudan (see fig. S12).

**East Timor** The only missing data point is the GDP growth rate for 1990. A report by the Australian government [44] suggests an average 10% economic growth rate from 1987–1997, due to heavy investment in construction. Given the data we already have, this implies GDP growth from 1987–1990 must have averaged 11.7%.

**Western Sahara, Saint Pierre Miquelon** No data are available for these territories. Western Sahara is under Moroccan control, and Saint Pierre Miquelon is a small French territory located near the Canadian province of Newfoundland. We thus simply use Moroccan and Canadian GDP per capita figures respectively.

**British overseas territories and Crown Dependencies** Four British territories have missing data: the Isle of Man, Gibraltar, the Falkland Islands, and Saint Helena, Ascension and Tristan da Cunha (in the following simply referred to as St. Helena), a small island territory in the Atlantic ocean. The Isle of Man government [45] provides the missing real GDP growth figures for 2022 and 2023, from which we can infer the missing real GDP/capita figures for those years. The government of Gibraltar [46] provides nominal data from 2009; we deflate using the British deflator as a proxy, and extrapolate back using the Isle of Man growth rate. Data for the Falklands real GDP is available from the government website, given at constant 2012 Falklands pounds (pegged to the British pound at an exchange rate of 1:1) [47]. We extrapolate back using the Isle of Man growth rates. Some historical data from 1980–1995 is available from the archives [48], which, when converted to real 2015 USD, agrees well with this extrapolation (see fig. S13). Lastly, nominal GDP data are provided by the St. Helena government website up to 2017. We use the British deflator to convert to real GDP and extrapolate back using the GDP growth rate for Brazil, a nearest neighbour with a similar GDP per capita.

**Faroe Islands** The World Bank provides real GDP/capita from 2008 onwards, and nominal GDP/capita from 1965 onwards, which we deflate using the Danish deflator to obtain a complete time series, see fig. S14.



**Figure S15:** Two tradeflow examples, showing total flow in constant 2015 USD. Shown are the BACI figures (solid blue line), as well as the Comtrade and IMF figures. Missing figures in the BACI dataset are extrapolated back using average growth rates (dotted lines).

## Bilateral trade

Trade flows are taken from the BACI harmonised tradeflow dataset [49, 50], and extrapolated back using the real growth rates calculated from the UN Comtrade and IMF Direction of trade statistics as given in the CEPII gravity dataset [51]. Since each dataset contains both origin- and destination-reported trade flows, we take the average growth rate (see fig. S15). Where values are missing, we simply set the trade flow to the last known value. If all values are NaN, we set the flow to 0.

## Total population, Life Expectancy, Birth and Death rates

We include the total population of the country of birth, country of origin, and country of destination for each edge  $(i, j, k)$ . These figures are taken from UN WPP dataset and scaled using a Yeo-Johnson transform (see fig. S16C–F).

## Distance

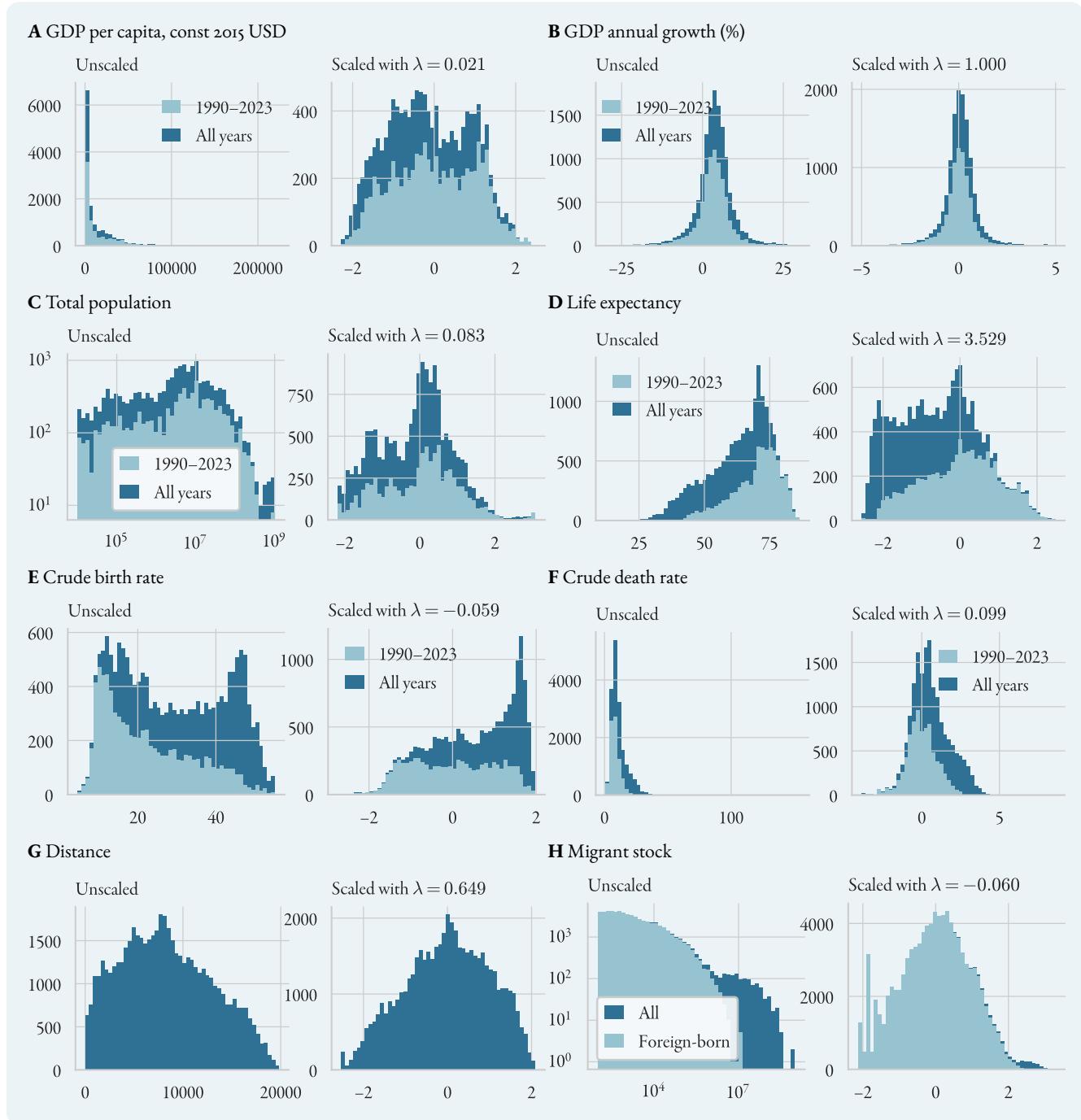
We use the `distw_harmonic` distance covariate from the CEPII dataset as a measure of distance. Since values for some years are missing, we take the average over the period 1990–2023. Values for the British Virgin Islands and the Isle of Man are missing; we approximate these by using the values for the neighbouring US Virgin Islands, and inserting a value of 80km for the distance US Virgin Islands–British Virgin Islands. For the Isle of Man, we use distance values for the United Kingdom and insert a value of 300km for the distance United Kingdom–Isle of Man. The covariate is then scaled using a Yeo-Johnson transform, see fig. S16G.

## Linguistic similarity

Linguistic similarity is taken the USITC Domestic and International Common Language Database [52]. We use the `cs1` (common spoken language) index. As detailed in the dataset description, values for 46 countries are missing. Following the authors, for these cases we use the average of the linguistic proximity `1pn` and common native language `cn1` instead.

## Religious similarity

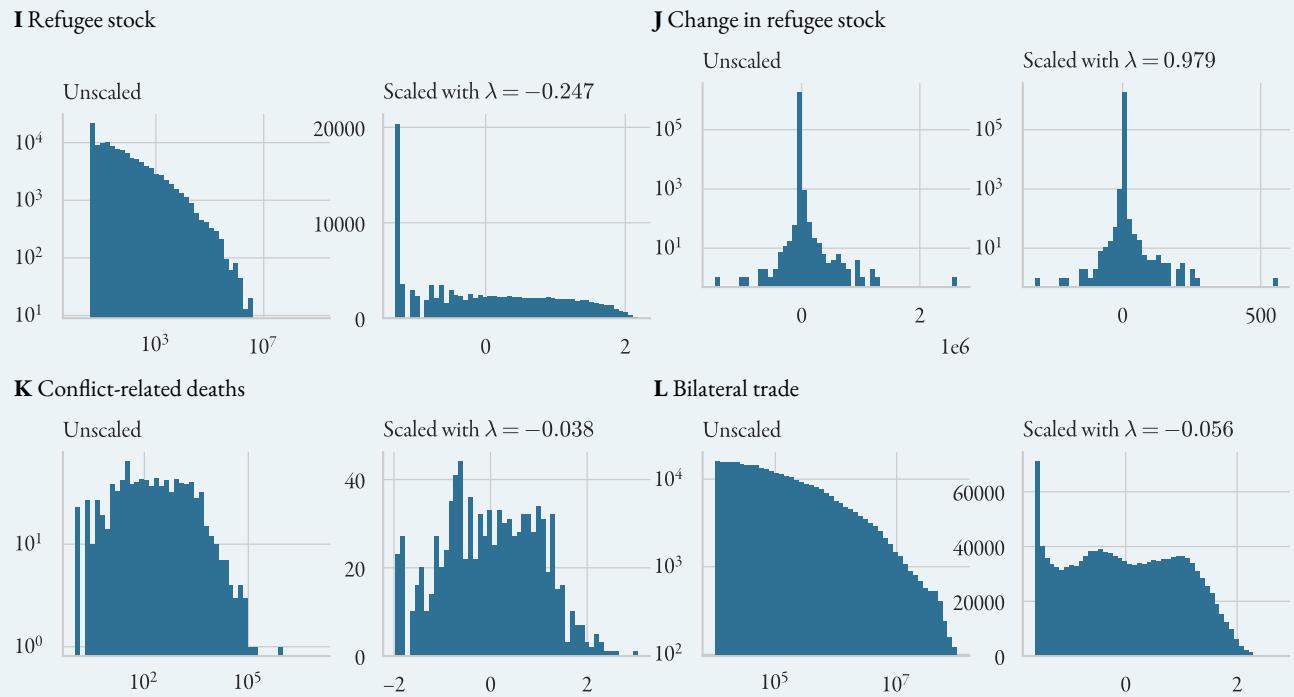
The Correlates of War database [53] details the religious composition  $\alpha$  of each country, that is the share of adherents to major religions in each country. Missing values are taken from the CIA World Factbook [54]. The religious similarity between two countries is then simply  $\langle \alpha_i, \alpha_j \rangle$ , where we exclude the ‘other’ (`othrgenpct`) category. This gives a



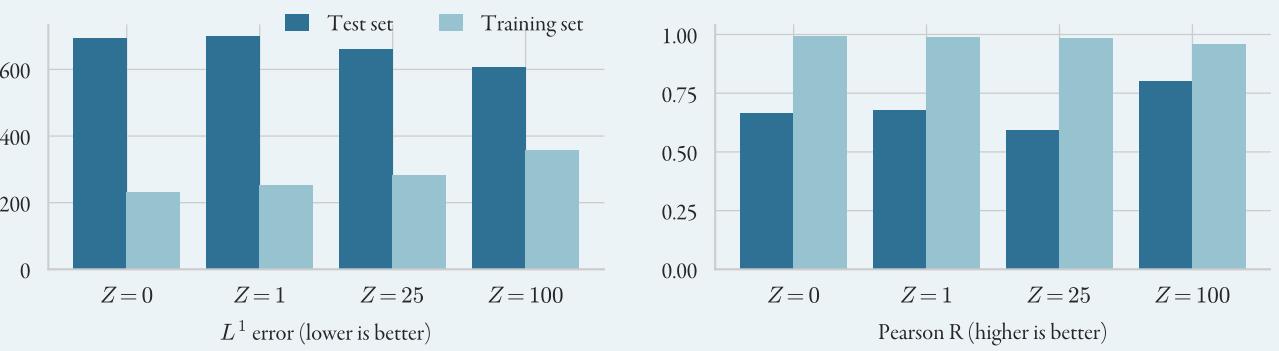
more accurate picture of religious proximity than the CEPII `comrelig` covariate, which only uses the share of Catholics, Protestants, and Muslims to construct a similarity score.

## EU Membership

This is taken from the CEPII gravity dataset. We construct two covariates,  $\text{EU}_{ij}$  and  $\text{EU}_{jk}$ , which are 1 if the indexed countries are both EU members, 0 else. As a binary covariate, it is left unscaled.



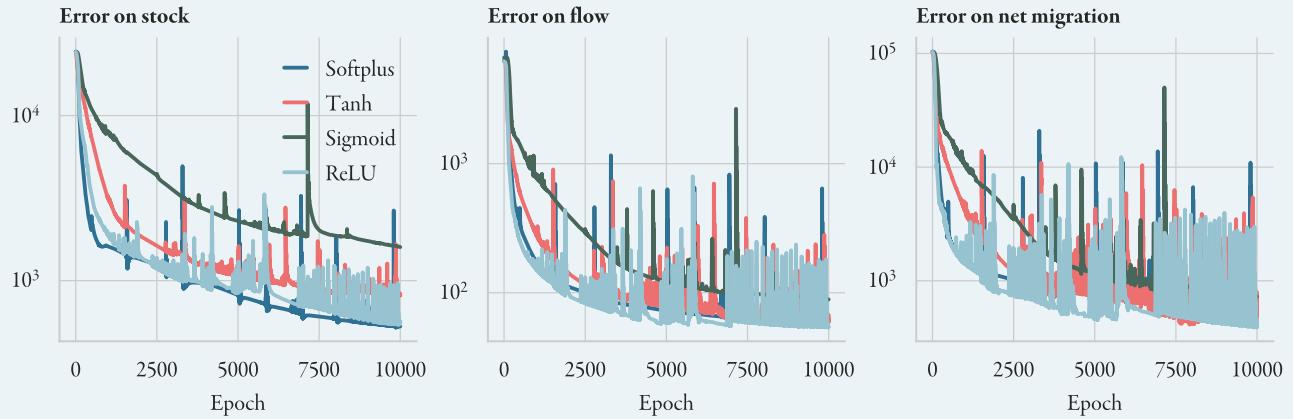
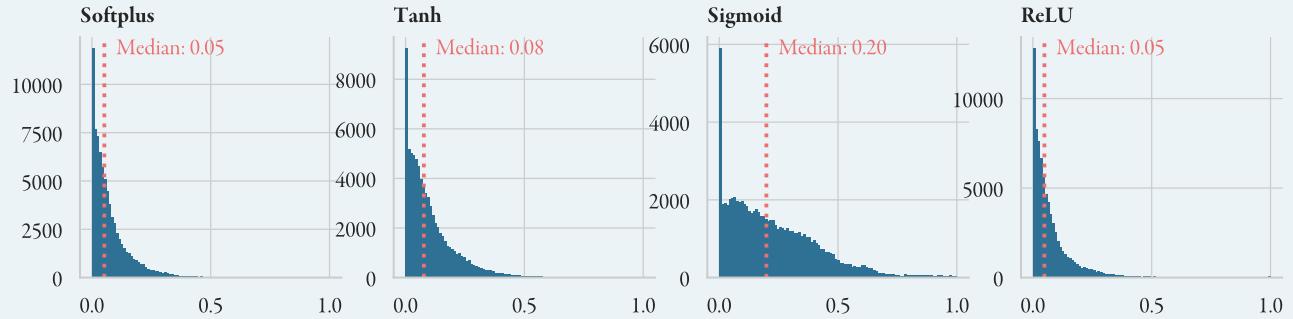
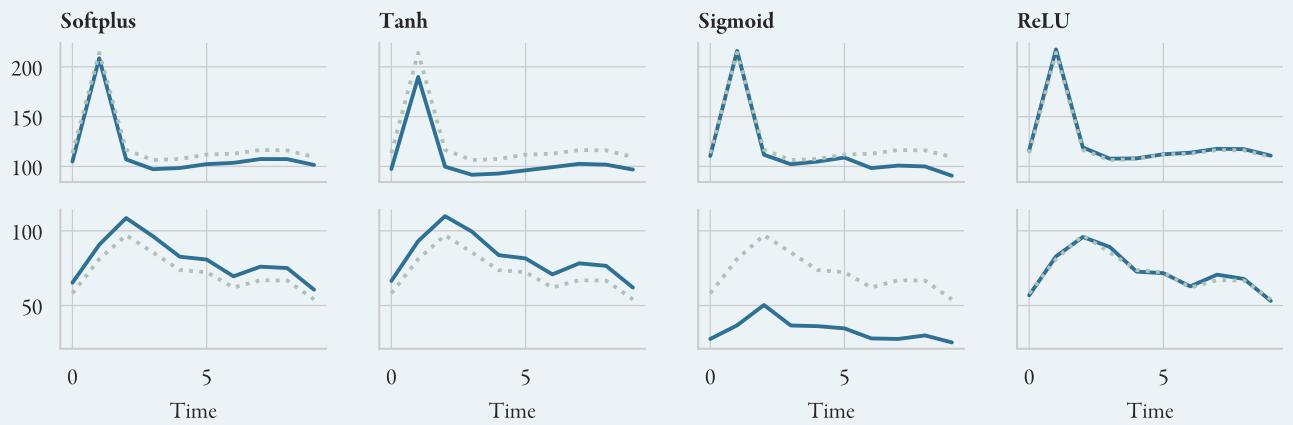
**Figure S16:** Scaling the input covariates. The transformation parameter  $\lambda$  is used to scale the raw data (left-hand figures) to an approximately normal distribution (right-hand figures) using the symmetrised Yeo-Johnson transform. The scaled values are then normalised to have zero mean and unit variance. The scaling  $\lambda$  is determined via maximum likelihood optimisation for the period 1990–2023.



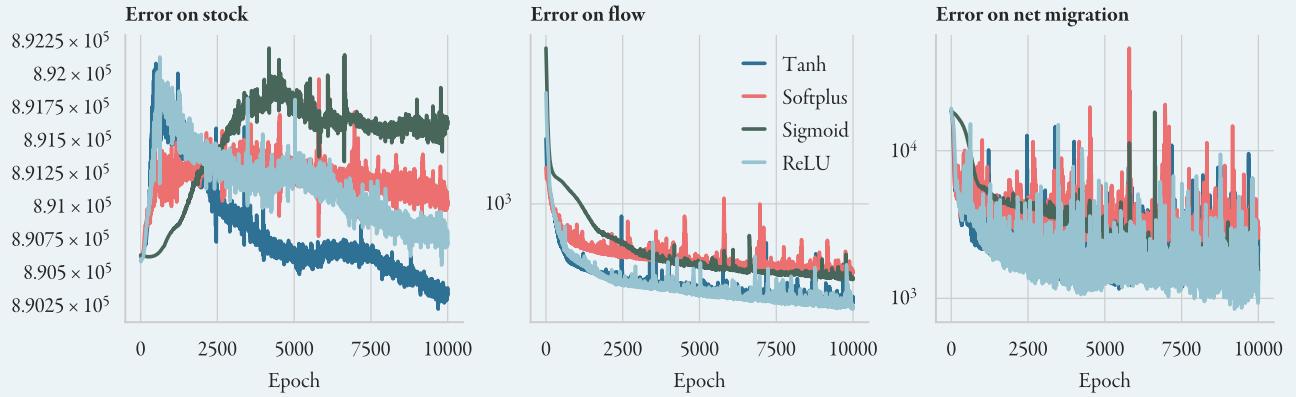
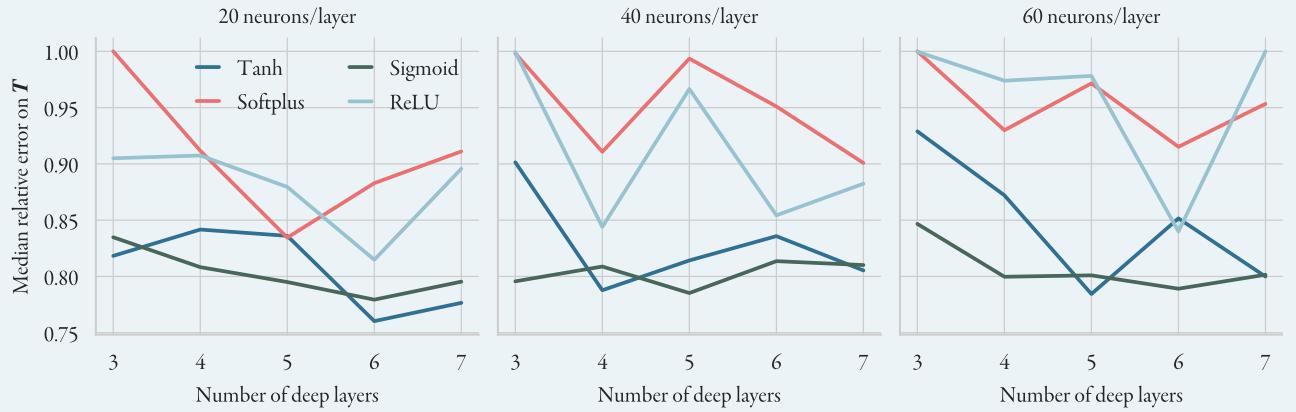
**Figure S17:** Adjusting the dimension  $Z$  of the latent state  $z$ .  $Z = 0$  indicates a regular feed-forward architecture with no recurrent structure. The far left, the average  $L^1$  error on the test and training set. Right: average correlation test and training set.

## Colonial ties

Colonial relations between two countries are taken from the USITC gravity dataset (`colony_of_destination_ever`). We include two covariates,  $\text{COL}_{jk}$  and  $\text{COL}_{ik}$ , which are 1 if  $j$  (or  $i$ ) was ever a colony of the destination  $k$ , and 0 else. As a binary covariate, it is left unscaled.

**A** Evolution of the  $L^1$  prediction errors during training**B** Relative error distribution of  $\hat{T}$ **C** Examples of inferred flow corridors  $T_{ijk}$ 

**Figure S18:** Validation of the neural network approach on noiseless, fully observed synthetic data. In all cases, we use a 3-layer neural network with 20 neurons per layer, but vary the activation function on the hidden layer. **A** The evolution of the error on the stocks, flows, and net migration as the models train. All models converge to the true target data, though at different rates. **B** The distribution of the relative error on the full table  $\hat{T}$  after 10,000 epochs. The median error is indicated. **C** Predicted and true (dashed) values  $T_{ijk}$  on two randomly selected edges. After 10,000 epochs (roughly 2 minutes of training on a GPU), the best-performing models have largely already converged to the true solution  $T$ .

**A** Evolution of the  $L^1$  prediction errors during training**B** Architecture tuning

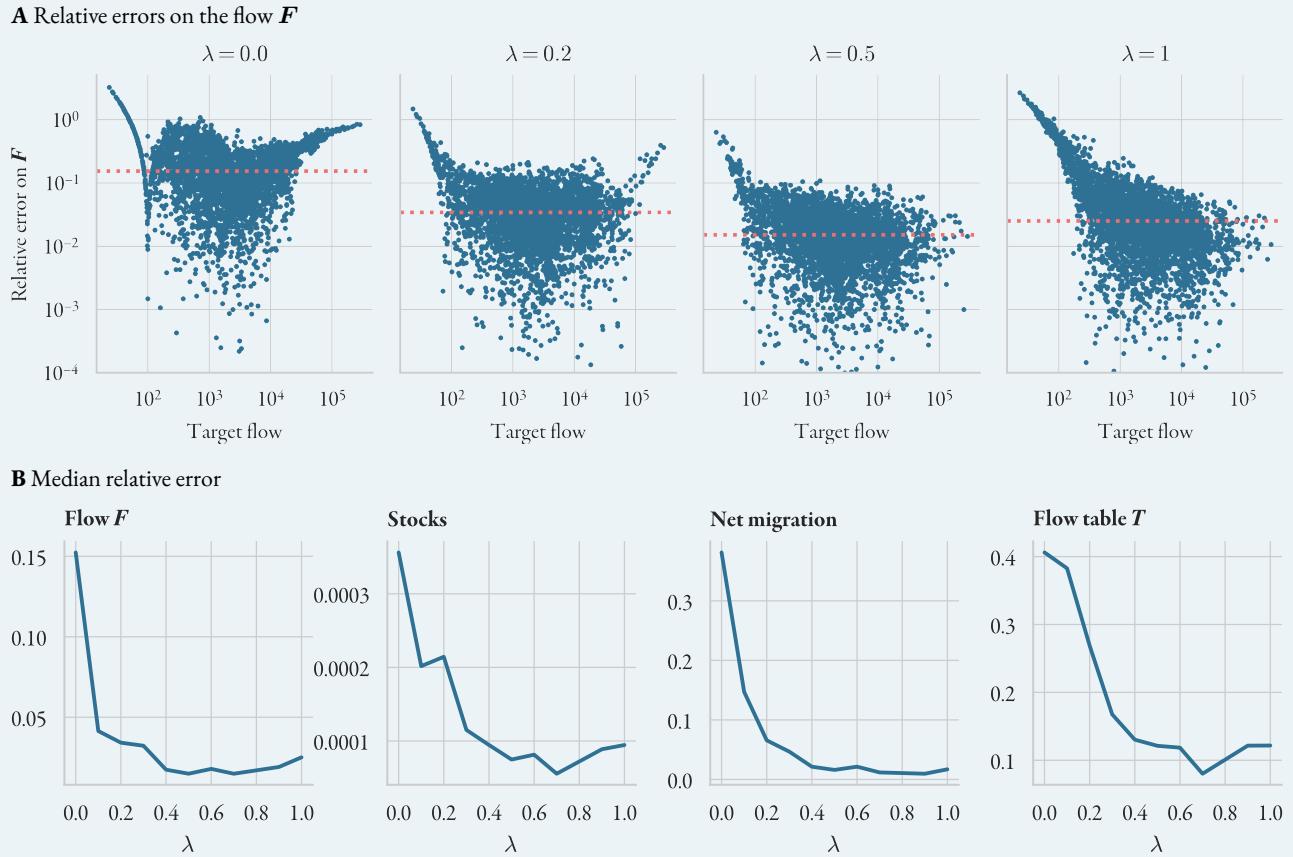
**Figure S19:** Tuning of the neural network hyperparameters on noisy, partially observed synthetic data. **A** The evolution of the error on the stocks, flows, and net migration is now noticeably slower and more jagged than in the noiseless case. **B** Comparison of the median relative error on  $\hat{\mathbf{T}}$  for various architectures.

## Validation on synthetic data

We validate our approach on a small synthetic dataset of migration flows between 30 destination and origin countries over 10 years. This is both to ensure our model is able to in principle infer the full flow table from observations of the stocks and flows, as well as to tune the hyperparameters. To do this, we select 30 random countries from the period 2010–2020, generate a random initial stock table  $\mathbf{S}(t_0)$ , and generate flows from

$$T_{ijk} = \eta \exp(\langle \chi_{ijk}, \alpha \rangle), \quad (6)$$

where  $\chi$  are the Yeo-Johnson-transformed covariate vectors,  $\eta = 100$  a scalar, and  $\alpha$  a vector of random coefficients drawn from an i.i.d uniform distribution,  $\alpha_i \sim \mathcal{U}[0, 0.5]$ . Note that the vector  $\chi$  contains the stocks  $S_{ij}$  and  $S_{ik}$  themselves, as in the full model; these are calculated recursively from  $\mathbf{T}$ . This produces a full synthetic dataset of stocks, flows, and net migration vectors on which we can validate the approach. We first assume all flow, net migration, and stock values are observed, and train the neural network for 10,000 epochs. We use a 3-layer neural network with 20 neurons in the hidden layers, and the same CeLU activation function on the final layer as in the main manuscript. We vary the choice of activation function in the hidden layer. As shown in fig. S18, the full flow table  $\mathbf{T}$  is then uniquely inferred from the data,



**Figure S2o:** Tuning the Yeo-Johnson scaling parameter  $\lambda$ . **A** A value of  $\lambda = 0.5$  balances the relative errors across different orders of magnitude of the target value. If  $\lambda$  is too small, all values are penalised equally, leading to poor performance across all orders of magnitude. A large value of  $\lambda$  leads to poor performance on small flows. The median relative error (red dotted line) is optimal for values around  $\lambda = 0.5$ . **B** Median relative error on all four datasets as a function of  $\lambda$ .

though accuracy varies with the choice of activation function. Naturally, since we are observing the full flow table  $\mathbf{F}$ , the net migration data are fully determined by  $\mathbf{F}$  and thus redundant.

To generate a more realistic dataset, we repeat the above process, adding 10% (multiplicative) noise to the stocks, 20% noise to flows, and 5% noise to the net migration data. We then randomly mask out 80% of flow corridors (that is, entire origin-destination corridors), 80% of net migration corridors, and 10% of individual stock values (not corridors). To tune the hyperparameters of the model, we perform a grid search on the depth, width, and activation functions of the network, and compare prediction accuracy on the true flow table  $\mathbf{T}$  (uncorrupted by noise) after 10,000 training epochs. Results are shown in fig. S19, indicating that the hyperbolic tangent activation function, together with a high number of layers, gives best prediction results.

To select the dimension of the latent space, we select 20% of flow corridors uniformly at random, and include only the remaining 80% in the training data; then we train an ensemble of neural networks with different latent space dimensions  $Z$  and compare performance. Fig. S17 shows the  $L^1$  error on the test set, the training loss, the average correlation on the test set, and the distribution of correlations on the flow corridors contained in the test set. As can be seen, a high latent space of dimension of 100 performs best of all the dimensions considered.

Finally, we set the Yeo-Johnson scaling parameters to a value of  $\lambda = 0.5$  for all target datasets, based on the following heuristic. The 0.5 value, intuitively, lies ‘half-way’ between the identity transform  $\lambda = 1$ , and the logarithmic transform

$\lambda = 0.0$ . In fig. S2o we show the learned flow values on the noiseless datasets for different values of the transformation parameters, suggesting that values between 0.5 and 0.7 are optimal. Small values of  $\lambda$  penalise all relative errors more equally, meaning a relative error of 50% on a flow of 100 contributes more to the loss than an error of 10% on a flow of 1 million—though in absolute terms these errors are of course orders of magnitude apart. For small  $\lambda$ , the model tries to optimise all targets equally well (or, in practice, equally badly). Large values of  $\lambda$ , conversely, mean that small flows are disregarded, and only the large values optimised. Centre values around  $\lambda = 0.5$  balances these considerations.

#### A Origin-destination flows

	DEMIG C2C				DEMIG TOTAL				Eurostat				IPUMS International				UN DESA IMFSC 2015				WPP 2024
Stock Diff.	0.49	0.51	0.58	0.52	0.27	0.53	0.71	0.65	0.58	0.58	0.35	0.06	0.69	0.61	0.47	0.50	0.47	0.33	0.04	0.69	0.74
Drop Negative	0.47	0.54	0.59	0.62	0.24	0.41	0.74	0.65	0.61	0.55	0.36	0.15	0.66	0.55	0.61	0.56	0.51	0.39	0.04	0.58	0.71
Stock Diff.	0.58	0.61	0.75	0.58	0.09	0.52	0.40	0.59	0.66	0.67	0.39	-0.07	0.43	0.51	0.62	0.62	0.74	0.46	-0.09	0.52	0.68
Reverse Negative	0.50	0.58	0.65	0.67	0.26	0.53	0.78	0.68	0.66	0.61	0.39	0.14	0.70	0.63	0.64	0.58	0.51	0.38	0.05	0.70	0.79
Migration rate	0.52	0.53	0.53	0.59	0.27	0.54	<b>0.81</b>	0.49	0.61	0.44	0.34	0.12	<b>0.75</b>	0.46	0.59	0.53	0.46	0.33	0.14	0.75	<b>1.00</b>
Dem. Acc.	0.69	0.78	0.80	0.74	0.37	0.57	0.81	0.68	0.84	0.74	<b>0.46</b>	0.33	0.74	0.59	0.73	0.78	0.76	0.47	0.31	0.74	1.00
Pseudo Bayes.	<b>0.70</b>	<b>0.86</b>	<b>0.84</b>	<b>0.77</b>	<b>0.45</b>	<b>0.60</b>	0.79	<b>0.83</b>	<b>0.89</b>	<b>0.80</b>	0.46	<b>0.57</b>	0.69	<b>0.66</b>	<b>0.80</b>	<b>0.84</b>	<b>0.77</b>	<b>0.56</b>	<b>0.54</b>	<b>0.77</b>	0.78
Neural																					
	Count	Log Count	Proportion	Migration Rate	Emigration Rate	Immigration Rate	Net Count	Count	Log Count	Proportion	Migration Rate	Emigration Rate	Immigration Rate	Count	Log Count	Proportion	Migration Rate	Emigration Rate	Immigration Rate	Net Count	

#### B Birth-destination flows

	UN CEPAL IMILA					DEMIG C2C					OECD					UN DESA IMFSC 2015						
Dem. Acc.	0.61	0.65	<b>0.94</b>	0.66	0.66	0.78	0.60	0.78	0.48	-0.03	0.52	0.58	0.59	0.59	0.34	0.16	0.65	0.81	0.69	0.90	-0.32	
Min. Open	<b>0.62</b>	0.64	0.94	<b>0.67</b>	0.67	0.78	0.60	0.79	0.48	0.20	0.53	0.58	0.59	0.59	0.34	0.17	0.65	0.80	0.69	0.90	-0.32	
Dem. Acc.	0.62	0.64	0.94	0.67	0.67	0.78	0.60	0.79	0.48	0.20	0.53	0.58	0.59	0.59	0.34	0.17	0.65	0.80	0.69	0.90	-0.32	
Min. Closed																						
Dem. Acc.	0.62	0.64	0.94	0.67	0.67	0.78	0.60	0.79	0.48	0.20	0.53	0.58	0.59	0.59	0.34	0.17	0.65	0.80	0.69	0.90	-0.32	
Pseudo Bayes.	0.55	<b>0.71</b>	0.89	0.57	<b>0.70</b>	<b>0.90</b>	<b>0.86</b>	<b>0.84</b>	<b>0.80</b>	<b>0.56</b>	<b>0.91</b>	<b>0.69</b>	<b>0.82</b>	<b>0.66</b>	<b>0.60</b>	<b>0.25</b>	<b>0.78</b>	<b>0.92</b>	<b>0.92</b>	<b>0.95</b>	<b>0.77</b>	<b>0.43</b>
Neural																						
	Count	Log Count	Proportion	Migration Rate	Immigration Rate	Count	Log Count	Proportion	Migration Rate	Emigration Rate	Immigration Rate	Count	Log Count	Proportion	Migration Rate	Emigration Rate	Immigration Rate	Count	Log Count	Proportion	Migration Rate	Immigration Rate

**Figure S21:** Comparison of various bilateral flow estimation techniques on unseen flow datasets. All values shown are Pearson R coefficients. Since all methods except our own only produce five-year flows, we aggregate the neural estimates up to five-year intervals for comparison purposes. The evaluation metrics are as in [55]. **A** Correlation on origin-destination flows. **B** Correlation on birth-destination flows.

## Comparison and validation

We further validate our approach on a dataset of unseen bilateral origin-destination and birth-destination flows. Since the datasets do not use a temporally consistent definition of migration (or include many different measures of migration), we do not calculate prediction errors directly, but rather use a series of correlation metrics similar to those presented in [55]. Unlike in fig. 6B in the main manuscript, we calculate correlations between the entire dataset  $\mathbf{Y}$ :

- *Count*: Pearson correlation coefficient on  $\mathbf{Y}$ ;
- *Log Count*: Pearson correlation coefficient on  $\log(\mathbf{Y} + 1)$ ;
- *Proportion*: Pearson correlation on  $y_{ij} / \sum_i y_{ij}$  if the observation  $y_{ij}$  was reported by the destination country, else  $y_{ij} / \sum_i y_{ij}$ ;
- *Migration rate*: Pearson correlation on  $y_{ij} / P_i$ , with  $P_i$  the total population of the origin;
- *Emigration rate*: Pearson correlation on  $\sum_i y_{ij} / P_i$ ;
- *Immigration rate*: Pearson correlation on  $\sum_i y_{ij} / P_j$ ;
- *Net count*: Pearson correlation on the net count  $\sum_i (y_{ij} - y_{ji})$ .

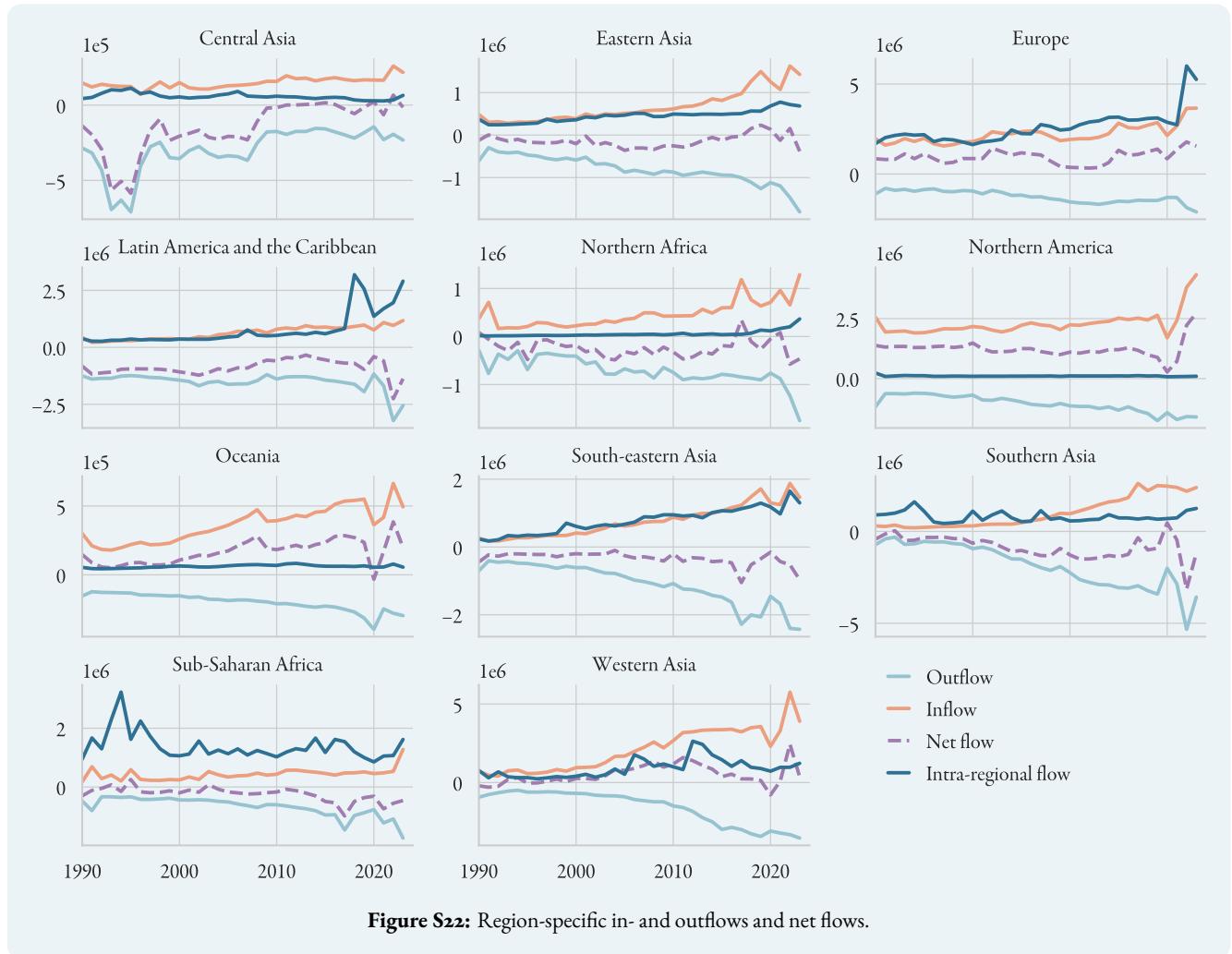
Note that in figure S21 we show correlations on both the total origin-destination flow ( $\sum_i T_{ijk}$ ) as well as the total birth-destination flow ( $\sum_j T_{ijk}$ ). We use the following validation datasets:

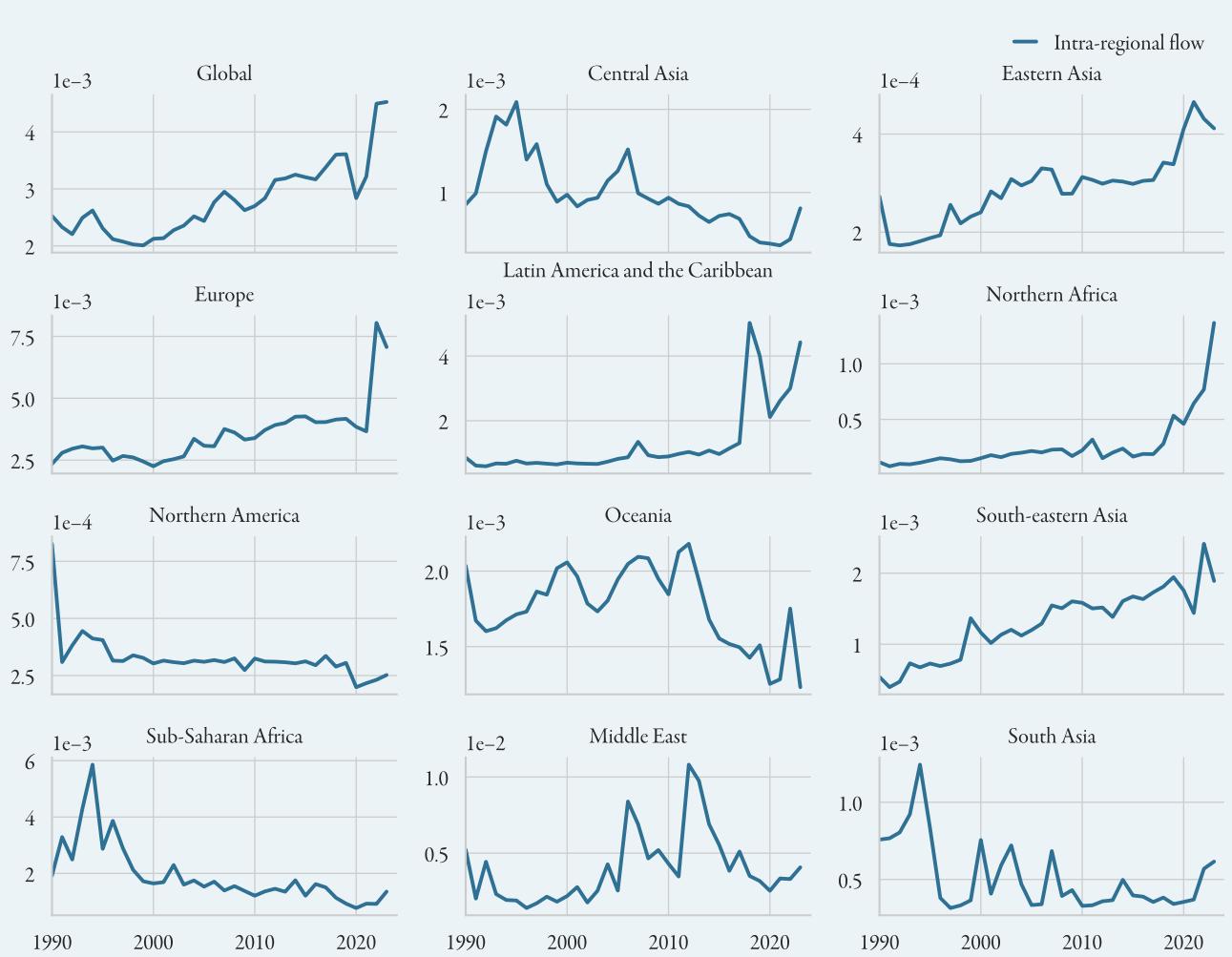
- DEMIG C2C [56]: bilateral flow data for 34 reporting (mostly European) countries from 1990–2011; this dataset contains both origin-destination and birth-destination flows.
- DEMIG TOTAL [57]: total immigration and emigration flows, as well as net counts, from 1990–2011.
- Eurostat [58]: bilateral origin-destination flows, mostly within Europe, from 1998–2019.
- IPUMS International [59]: immigration totals from census data covering the period 1990–2016.
- UN DESA IMFSC [60]: bilateral origin-destination and birth-destination flows, reported by 45 (mainly European) countries from 1990–2013.
- UN CEPAL IMILA [61]: bilateral birth-destination flow data to and from Latin American countries from 1990–2013. Excludes return migration of native-born emigrants.
- OECD [62]: birth-destination flows for OECD countries, 1995–2013.
- WPP [33]: UN WPP net migration estimates for all countries, 2024 revision, 1990–2020.

Figure S21 shows the metrics for the various stock differencing methods outlined in the introduction, as well as our neural estimates. Since all methods except our own produce five-year flows, we aggregate our results up to the five-year level. The closed demographic accounting methods have been adjusted to match the same demographic residuals used to calculate the UN WPP net migration figures, hence their correlation with that dataset is 1.

## Additional figures

Here we show some additional plots for on regional flows. Figure S22 shows the outflow, inflow, net flow, and intra-regional flow for all regions, while figure S23 shows the intra-regional flows per capita. Also shown is the per capita total global migration flow: as is visible, the same trend of rising global migration emerges, even when accounting for the global population increase.





**Figure S23:** Per-capita intra-regional flows. Also shown are the total global flows (top left). The trend matches that shown in fig. 2A in the main manuscript.

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