

# Supplementary Information for Learning Risk Preferences Through Social Interaction: An Active Inference Approach

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## ABSTRACT

This supplementary file provides additional details for the manuscript titled "Learning Risk Preferences Through Social Interaction: An Active Inference Approach". It includes three appendices containing mathematical derivations for equations presented in the main text, as well as a supplementary figure.

## Appendix S1: Simplifying the Variational energy

It is possible to simplify the expression  $\ln [1 - S(F(\rho^{(k)}))]$  in the following manner:

$$\begin{aligned}\ln [1 - S(F(\rho^{(k)}))] &= \ln \left[ 1 - \frac{1}{1 + \exp(-\beta \times F(\rho^{(k)}))} \right] \\ &= \ln \left[ \frac{\exp(-\beta \times F(\rho^{(k)}))}{1 + \exp(-\beta \times F(\rho^{(k)}))} \right] \\ &= \ln [\exp(-\beta \times F(\rho^{(k)}))] + \ln \left[ \frac{1}{1 + \exp(-\beta \times F(\rho^{(k)}))} \right] \\ &= -\beta \times F(\rho^{(k)}) + \ln [S(F(\rho^{(k)}))]\end{aligned}\tag{S1.1}$$

Substituting the value of  $\ln [1 - S(F(\rho^{(k)}))]$  in Variational energy formula and carrying out the necessary operations yields the following expression:

$$\begin{aligned}I(\rho^{(k)}) &= C^{(k)} \times \ln [S(F(\rho^{(k)}))] + (1 - C^{(k)}) \times (-\beta \times F(\rho^{(k)}) + \ln [S(F(\rho^{(k)}))]) \\ &\quad - \frac{(\rho^{(k)} - \mu_p^{(k-1)})^2}{2\sigma_p^{(k-1)}} \\ &= \ln [S(F(\rho^{(k)}))] + \beta \times (C^{(k)} - 1) \times F(\rho^{(k)}) - \frac{(\rho^{(k)} - \mu_p^{(k-1)})^2}{2\sigma_p^{(k-1)}}\end{aligned}\tag{S1.2}$$

## Appendix S2: Obtaining a quadratic approximation of the Variational energy $I(\rho^{(k)})$

Here one can observe the second order approximation of the Variational energy, resulting in an update equation. The equation obtained for  $I(\rho^{(k)})$  contains non-quadratic terms:

$$I(\rho^{(k)}) = \ln \left[ S \left( F(\rho^{(k)}) \right) \right] + \beta \times (C_0^{(k)} - 1) \times F(\rho^{(k)}) - \frac{(\rho^{(k)} - \mu_p^{(k-1)})^2}{2\sigma_p^{(k-1)}}$$

In order to approximate a Gaussian distribution, it is necessary to find a quadratic polynomial  $\tilde{I}(\rho^{(k)})$  that approximates  $I(\rho^{(k)})$ . By doing so, the Gaussian distribution can be accurately represented.

$$\tilde{I}(\rho^{(k)}) = -\frac{(\rho^{(k)} - \mu_p^{(k)})^2}{2\sigma_p^{(k)}} \quad (\text{S2.1})$$

When the second order Taylor series of function  $I(\rho^{(k)})$  is expanded around a particular expansion point, an approximate quadratic function  $\tilde{I}(\rho^{(k)})$  can be obtained. This technique is referred to as Laplace approximation when the Taylor expansion is carried out at the peak of  $I(\rho^{(k)})$ <sup>1</sup>. An alternative site for the expansion point would be the mean value of  $\rho^{(k)}$  that has been previously updated, symbolized by  $\mu_p^{(k-1)}$ <sup>2</sup>. The second suggestion is what we are going to use here.

Considering  $\tilde{I}(\rho^{(k)})$  to be the second order Taylor expansion of  $I(\rho^{(k)})$  about  $\mu_p^{(k-1)}$ , the first and second derivatives of both  $\tilde{I}(\rho^{(k)})$  and  $I(\rho^{(k)})$  are equivalent at this point:

$$\begin{aligned} \left( \frac{\partial I}{\partial \rho^{(k)}} \right) \bigg|_{\rho^{(k)} = \mu_p^{(k-1)}} &= \left( \frac{\partial \tilde{I}}{\partial \rho^{(k)}} \right) \bigg|_{\rho^{(k)} = \mu_p^{(k-1)}} \\ \left( \frac{\partial^2 I}{\partial (\rho^{(k)})^2} \right) \bigg|_{\rho^{(k)} = \mu_p^{(k-1)}} &= \left( \frac{\partial^2 \tilde{I}}{\partial (\rho^{(k)})^2} \right) \bigg|_{\rho^{(k)} = \mu_p^{(k-1)}} \end{aligned} \quad (\text{S2.2})$$

The derivatives of  $\tilde{I}(\rho^{(k)})$  at the point  $\mu_p^{(k-1)}$  can be determined from Eq (S2.1). Substituting these values into Eq(S2.2) leads to the following results:

$$\begin{aligned} \left( \frac{\partial I}{\partial \rho^{(k)}} \right) \bigg|_{\rho^{(k)} = \mu_p^{(k-1)}} &= \frac{\mu_p^{(k)} - \mu_p^{(k-1)}}{\sigma_p^{(k)}} \\ \left( \frac{\partial^2 I}{\partial (\rho^{(k)})^2} \right) \bigg|_{\rho^{(k)} = \mu_p^{(k-1)}} &= -\frac{1}{\sigma_p^{(k)}} \end{aligned} \quad (\text{S2.3})$$

Rearranging Eq(S2.3) yields into an updating equation for  $\mu_p^{(k)}$  and  $\sigma_p^{(k)}$ :

$$\begin{aligned} \mu_p^{(k)} &= \mu_p^{(k-1)} + \sigma_p^{(k)} \times \left( \frac{\partial I}{\partial \rho^{(k)}} \right) \bigg|_{\rho^{(k)} = \mu_p^{(k-1)}} \\ \frac{1}{\sigma_p^{(k)}} &= -\left( \frac{\partial^2 I}{\partial (\rho^{(k)})^2} \right) \bigg|_{\rho^{(k)} = \mu_p^{(k-1)}} \end{aligned} \quad (\text{S2.4})$$

The final step to obtain the explicit update equations for  $\mu_p^{(k)}$  and  $\sigma_p^{(k)}$  is to calculate the first and second derivatives of

$I(\rho^{(k)})$  with respect to  $\mu_p^{(k-1)}$ . We begin by computing the first derivative.

$$\begin{aligned}
\frac{\partial I}{\partial \rho^{(k)}} &= \frac{\partial}{\partial \rho^{(k)}} \left\{ \ln S(F(\rho^{(k)})) + \beta \times (C_o^{(k)} - 1) \times F(\rho^{(k)}) - \frac{(\rho^{(k)} - \mu_p^{(k-1)})^2}{2\sigma_p^{(k-1)}} \right\} \\
&= \frac{\partial}{\partial \rho^{(k)}} \left\{ \ln S(F(\rho^{(k)})) + \beta \times (C_o^{(k)} - 1) \times F(\rho^{(k)}) \right\} - \frac{\partial}{\partial \rho^{(k)}} \left\{ \frac{(\rho^{(k)} - \mu_p^{(k-1)})^2}{2\sigma_p^{(k-1)}} \right\} \\
&= \frac{\partial}{\partial \rho^{(k)}} \left\{ \ln S(F(\rho^{(k)})) \right\} + \beta \times (C_o^{(k)} - 1) \times \frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} - \frac{\rho^{(k)} - \mu_p^{(k-1)}}{\sigma_p^{(k-1)}}
\end{aligned} \tag{S2.5}$$

Using a chain rule, it is possible to derive  $\frac{\partial}{\partial \rho^{(k)}} \left\{ \ln [S(F(\rho^{(k)}))] \right\}$  as follows:

$$\frac{\partial}{\partial \rho^{(k)}} \left\{ \ln [S(F(\rho^{(k)}))] \right\} = \frac{1}{S(F(\rho^{(k)}))} \times \frac{\partial S}{\partial F(\rho^{(k)})} \times \frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} \tag{S2.6}$$

where  $\frac{\partial S}{\partial F(\rho^{(k)})}$  is:

$$\frac{\partial S}{\partial F(\rho^{(k)})} = \beta \times [S(F(\rho^{(k)}))] \times [1 - S(F(\rho^{(k)}))] \tag{S2.7}$$

Finally we have:

$$\frac{\partial}{\partial \rho^{(k)}} \left\{ \ln [S(F(\rho^{(k)}))] \right\} = \beta \times \frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} \times [1 - S(F(\rho^{(k)}))] \tag{S2.8}$$

The following equation for  $\frac{\partial I}{\partial \rho^{(k)}}$  can be calculated by substituting the value achieved for  $\frac{\partial}{\partial \rho^{(k)}} \left\{ \ln [S(F(\rho^{(k)}))] \right\}$  in the Eq (S2.5).

$$\begin{aligned}
\frac{\partial I}{\partial \rho^{(k)}} &= \beta \times \frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} [1 - S(F(\rho^{(k)}))] + \beta \times \frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} \times (C_o^k - 1) - \frac{\rho^{(k)} - \mu_p^{(k-1)}}{\sigma_p^{(k-1)}} \\
&= \beta \times \frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} \times [C_o^k - S(F(\rho^{(k)}))] - \frac{\rho^{(k)} - \mu_p^{(k-1)}}{\sigma_p^{(k-1)}}
\end{aligned} \tag{S2.9}$$

By differentiating Eq (S2.9), we can obtain the expression for  $\frac{\partial^2 I}{\partial (\rho^{(k)})^2}$ :

$$\begin{aligned}
\frac{\partial^2 I}{\partial (\rho^{(k)})^2} &= \frac{\partial}{\partial \rho^{(k)}} \left( \frac{\partial I}{\partial \rho^{(k)}} \right) \\
&= \frac{\partial}{\partial \rho^{(k)}} \left( \beta \times \frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} \times [C_o^k - S(F(\rho^{(k)}))] - \frac{\rho^{(k)} - \mu_p^{(k-1)}}{\sigma_p^{(k-1)}} \right) \\
&= \frac{\partial}{\partial \rho^{(k)}} \left( \beta \times \frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} \times [C_o^k - S(F(\rho^{(k)}))] \right) - \frac{\partial}{\partial \rho^{(k)}} \left( \frac{\rho^{(k)} - \mu_p^{(k-1)}}{\sigma_p^{(k-1)}} \right) \\
&= \beta \times \frac{\partial^2 F(\rho^{(k)})}{\partial (\rho^{(k)})^2} \times [C_o^k - S(F(\rho^{(k)}))] - \beta \times \left( \frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} \right)^2 \times \frac{\partial S}{\partial F(\rho^{(k)})} - \frac{1}{\sigma_p^{(k-1)}}
\end{aligned} \tag{S2.10}$$

Omitting the initial term which incorporates the second order derivative  $\frac{\partial^2 F(\rho^{(k)})}{\partial (\rho^{(k)})^2}$  yields a more straightforward formula:

$$\frac{\partial^2 I}{\partial (\rho^{(k)})^2} \cong -\beta^2 \times \left( \frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} \right)^2 \times [S(F(\rho^{(k)}))] \times [1 - S(F(\rho^{(k)}))] - \frac{1}{\sigma_p^{(k-1)}} \tag{S2.11}$$

The value of  $\frac{\partial^2 I}{\partial(\rho^{(k)})^2}$  is always kept at a negative level due to the presence of negative components. This ensures that  $\sigma_\rho^{(k)}$  is maintained at a positive value, which is a necessary requirement based on its underlying characteristics.

Lastly, by substituting the values of  $\frac{\partial I}{\partial \rho^{(k)}}$  and  $\frac{\partial^2 I}{\partial(\rho^{(k)})^2}$  at the point  $\rho^{(k)} = \mu_\rho^{(k-1)}$  into Equation Eq(S2.4), the update equation for  $\mu_\rho^{(k)}$  and  $\sigma_\rho^{(k)}$  can be obtained:

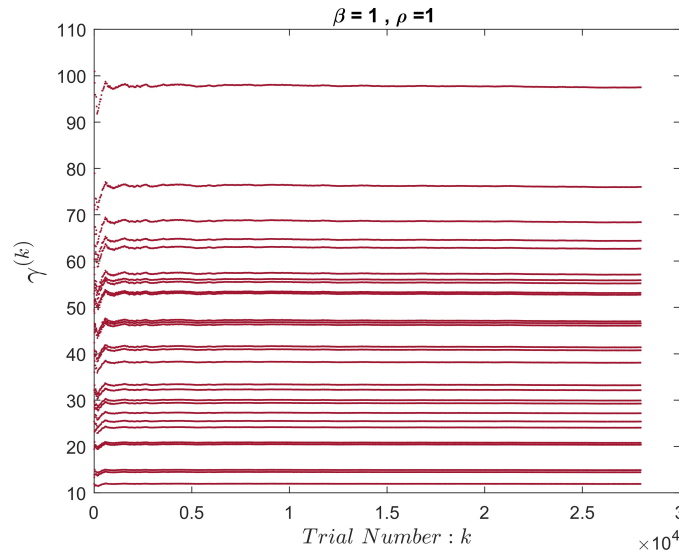
$$\begin{aligned}\mu_\rho^{(k)} &= \mu_\rho^{(k-1)} + \beta \times \left( \frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} \bigg|_{\rho^{(k)}=\mu_\rho^{(k-1)}} \right) \times \sigma_\rho^{(k)} \times \left( C_0^{(k)} - S\left(F(\mu_\rho^{(k-1)})\right) \right) \\ \frac{1}{\sigma_\rho^{(k)}} &= \frac{1}{\sigma_\rho^{(k-1)}} + \beta^2 \times \left( \frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} \bigg|_{\rho^{(k)}=\mu_\rho^{(k-1)}} \right)^2 \times \left[ S\left(F(\mu_\rho^{(k-1)})\right) \right] \times \left[ 1 - S\left(F(\mu_\rho^{(k-1)})\right) \right]\end{aligned}\quad (S2.12)$$

Where the value of  $\frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} \bigg|_{\rho^{(k)}=\mu_\rho^{(k-1)}}$  is:

$$\begin{aligned}\frac{\partial F(\rho^{(k)})}{\partial \rho^{(k)}} \bigg|_{\rho^{(k)}=\mu_\rho^{(k-1)}} &= \frac{\partial}{\partial \rho^{(k)}} \left\{ p^k \times (r^k)^{\rho^k} - U_s \right\} \bigg|_{\rho^{(k)}=\mu_\rho^{(k-1)}} \\ &= p^k \times \ln r^k \times (r^k)^{\mu_\rho^{(k-1)}}\end{aligned}\quad (S2.13)$$

See method section in the main manuscript for the definition of function  $F(\cdot)$ .

## Appendix S3: Supplementary Figure 1



**Supplementary Figure 1. Variation of  $\gamma^{(k)}$  across trials.** As detailed in the main text,  $\gamma^{(k)}$  is calculated as the derivative of  $F(\rho^{(k)})$  with respect to  $\rho^{(k)}$ . The value of  $\gamma^{(k)}$  is a function of the reward amount and the probability of the proposed gamble for each trial.

## References

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2. Mathys, C., Daunizeau, J., Friston, K. J. & Stephan, K. E. A bayesian foundation for individual learning under uncertainty. *Frontiers in human neuroscience* **5**, 39 (2011).