

Online Supplement to: Efficient Modeling of Quasi-Periodic Data with Seasonal Gaussian Process

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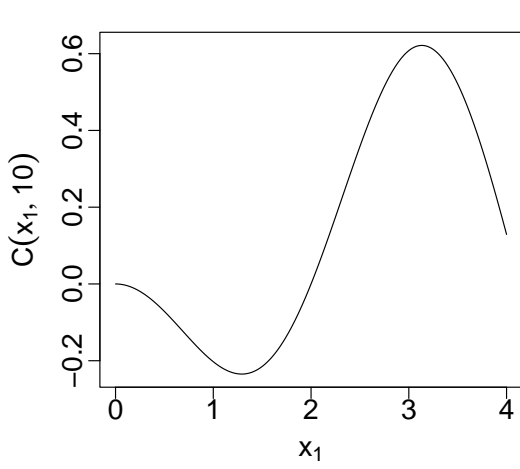
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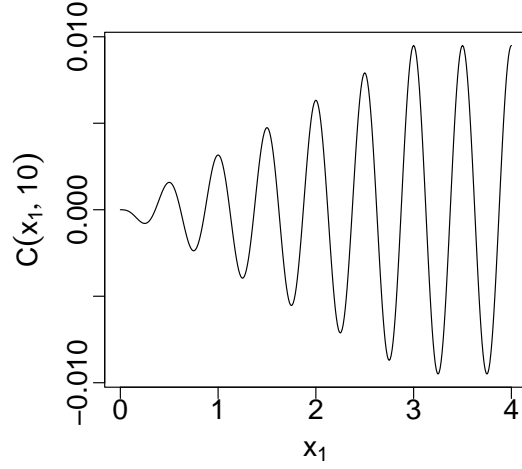
S1. Additional figures and table

	0.025 Quantile	Median	0.975 Quantile
$\sigma_{tr}(10)$	0.040 (15.336)	0.960 (37.655)	9.897 (106.646)
$\sigma_1(10)$	0.126 (–)	0.270 (–)	0.509 (–)
$\sigma_{\frac{1}{2}}(10)$	0.004 (–)	0.054 (–)	0.146 (–)
$\sigma_{\frac{44}{12}}(10)$	0.162 (–)	0.518 (–)	0.889 (–)
$\sigma_{9.1}(10)$	0.025 (–)	0.391 (–)	1.429 (–)
$\sigma_{10.4}(10)$	0.022 (–)	0.430 (–)	1.568 (–)
σ_{ϵ}	0.553 (0.582)	0.584 (0.615)	0.620 (0.651)

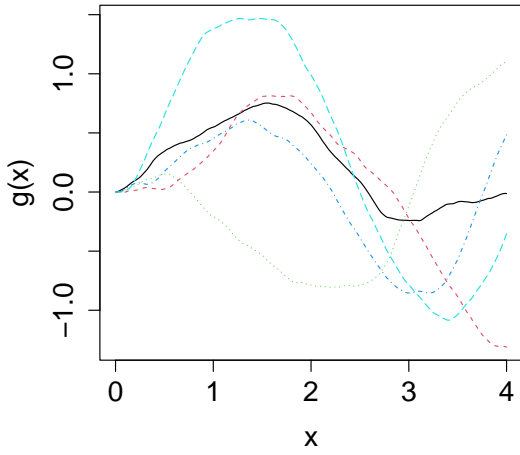
Table S1: Posterior summary of standard deviation parameters for the CO2 example in Section 5.4. Results from M2 are shown in parenthesis.



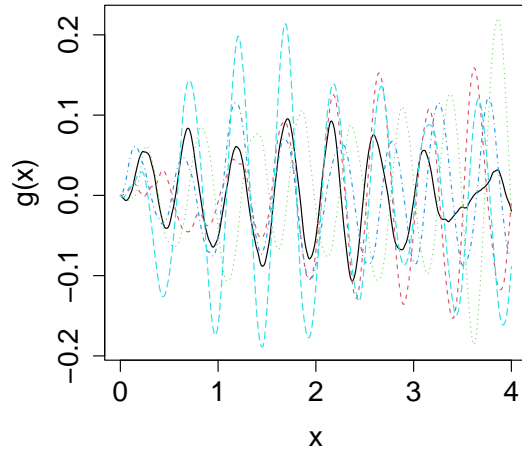
(a) $\alpha = \pi/2$: Covariance



(b) $\alpha = 4\pi$: Covariance

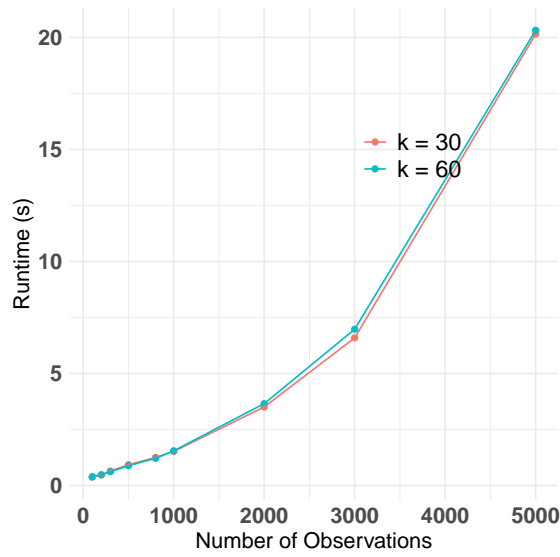


(c) $\alpha = \pi/2$: Samples

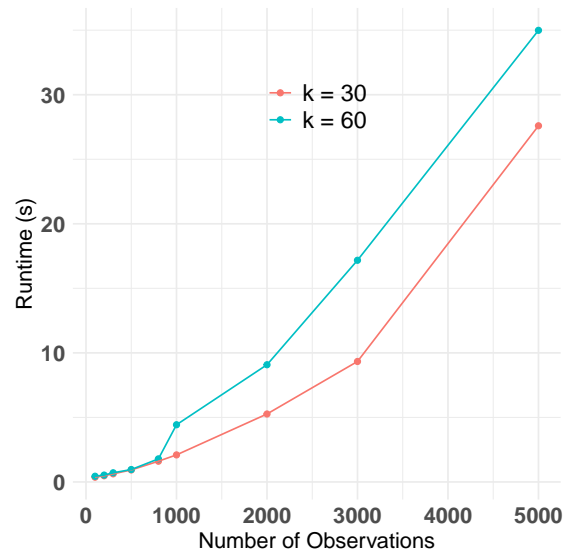


(d) $\alpha = 4\pi$: Samples

Figure S1: Figures (a,b) show covariance functions of the sGP with different α , where the first argument in the covariance function is fixed at 3. Figures (c,d) display five sample paths from the two sGPs. The frequency parameter α equals to $\pi/2$ in (a,c) and 4π in (b,d), and the SD parameter $\sigma = 1$ in both sGPs.

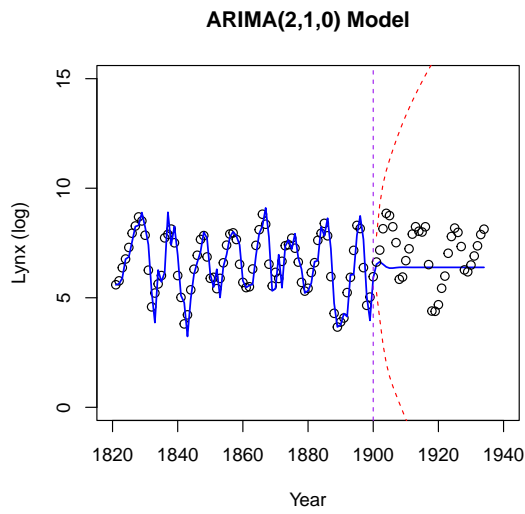


(a) Easy Setting (A)

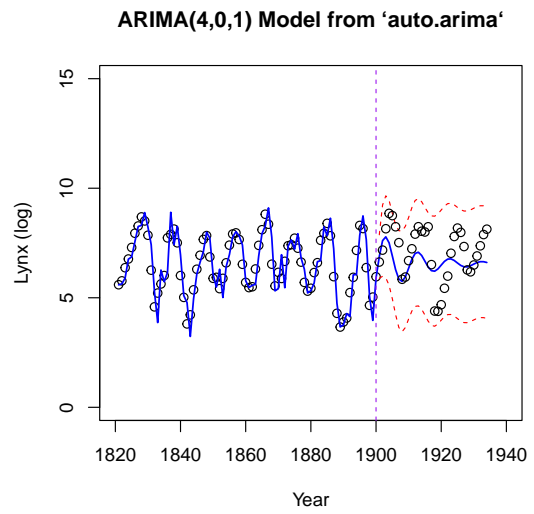


(b) Hard Setting (B)

Figure S2: Average runtimes using the sB-spline approximation with $k = 30$ (red) and $k = 60$ (blue) for the two simulation settings in Section 4.1, when the sample size n varies. Each average is computed from ten replications.

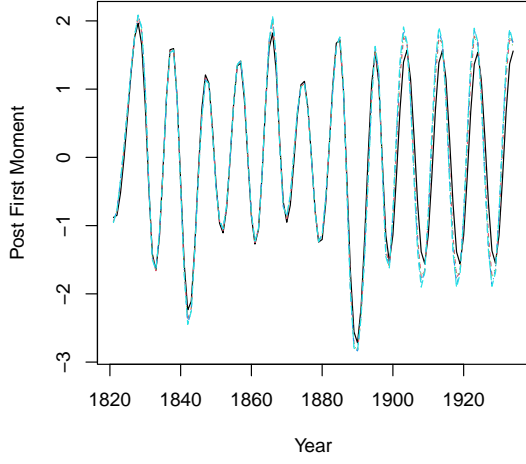


(a)

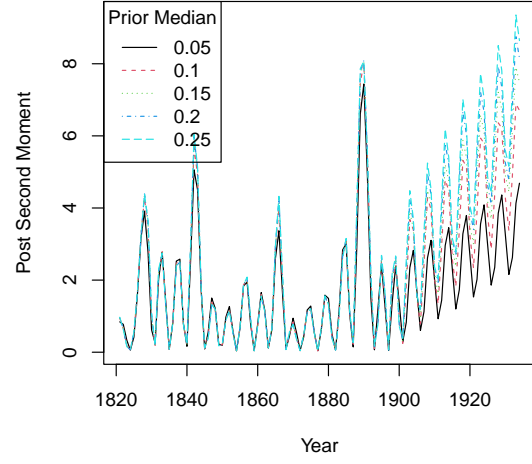


(b)

Figure S3: Additional comparisons for Section 4.3. Frequentist ARIMA models fitted using maximum likelihood estimation: (a) ARIMA with order fixed at (2,1,0); (b) ARIMA with optimal order (4,0,1) selected based on AIC.



(a) First Moment



(b) Second Moment

Figure S4: Additional sensitivity analysis for Section 4.3. Posterior moments obtained from the sGP model, with the median in the PSD prior varying over a range of possible values.

S2. Derivation of the sGP covariance

Proposition S1 (Covariance Function of the Seasonal Gaussian Process). *Let $g \sim sGP_\alpha(\sigma)$. Then g has a covariance function:*

$$\begin{aligned} C(x_1, x_2) &= \left(\frac{\sigma}{\alpha}\right)^2 \left[\frac{x_1}{2} \cos(\alpha(x_2 - x_1)) - \frac{\cos(\alpha x_2) \sin(\alpha x_1)}{2\alpha} \right] \\ &= \left(\frac{\sigma}{\alpha}\right)^2 \left[\frac{\cos(\alpha x_2) x_1}{2} \cos(\alpha x_1) + \left(\frac{\sin(\alpha x_2) x_1}{2} - \frac{\cos(\alpha x_2)}{2\alpha} \right) \sin(\alpha x_1) \right], \end{aligned} \quad (1)$$

for any $x_1, x_2 \in \Omega$ such that $x_1 \leq x_2$.

Proof. It is obvious that the differential operator L is linear. Define $\mathbf{g}_{aug}(x) = (g(x), g'(x))^T$ and therefore $\mathbf{g}'_{aug}(x) = (g'(x), g''(x))^T$, then the SDE can be rewritten in the vector form:

$$\mathbf{g}'_{aug} = \mathbf{F} \mathbf{g}_{aug} + \mathbf{J} W, \quad (2)$$

where $\mathbf{F} = \begin{bmatrix} 0 & 1 \\ -\alpha^2 & 0 \end{bmatrix}$ and $\mathbf{J} = \begin{bmatrix} 0 \\ \sigma \end{bmatrix}$.

Using the result from Särkkä and Solin (2019) (section 4.3), the solution of the linear SDE can be written as:

$$\begin{aligned} \mathbf{g}_{aug}(x) &= \exp(\mathbf{F}x) \mathbf{g}_{aug}(0) + \int_0^x \exp(\mathbf{F}(x - \tau)) \mathbf{J} W(\tau) d\tau \\ &= \int_0^x \exp(\mathbf{F}(x - \tau)) \mathbf{J} W(\tau) d\tau, \end{aligned} \quad (3)$$

where $\exp(\mathbf{F}x)$ denotes the matrix exponential defined as $\exp(\mathbf{F}x) = \sum_k \frac{\mathbf{F}^k x^k}{k!}$.

Note that $\mathbf{F}^{2k} = (-\alpha^2)^k \mathbf{I}$ and $\mathbf{F}^{2k+1} = (-\alpha^2)^k \mathbf{F}$. With Taylor series, the first component of $\mathbf{g}_{aug}(x)$ can be therefore written as:

$$g(x) = \int_0^x \frac{\sigma}{\alpha} \sin(\alpha(x - \tau)) W(\tau) d\tau. \quad (4)$$

Assume arbitrary $0 < x_1 \leq x_2$, the covariance function can be computed for g as:

$$\begin{aligned} C(x_1, x_2) &= \int_0^{x_1} \frac{\sigma}{\alpha} \sin(\alpha(x_1 - \tau)) \frac{\sigma}{\alpha} \sin(\alpha(x_2 - \tau)) d\tau \\ &= \left(\frac{\sigma}{\alpha}\right)^2 \left[\frac{x_1}{2} \cos(\alpha(x_2 - x_1)) - \frac{\cos(\alpha x_2) \sin(\alpha x_1)}{2\alpha} \right], \end{aligned} \quad (5)$$

using properties of Gaussian white noise (Harvey, 1990).

□

S3. Proof of the State-Space Representation

Theorem S1 (State Space Representation of the sGP). *Consider $g \sim sGP_\alpha(\sigma)$, and let $\mathbf{s} = \{s_1, \dots, s_n\} \subset \Omega$ be sorted with spacing $d_1 = s_1$ and $d_i = s_i - s_{i-1}$ for $i \in \{2, \dots, n\}$. Then $\mathbf{g}_{aug}(s_i) = [g(s_i), g'(s_i)]^T$ can be written as a Markov model:*

$$\mathbf{g}_{aug}(s_{i+1}) = \mathbf{R}_{i+1} \mathbf{g}_{aug}(s_i) + \boldsymbol{\epsilon}_{i+1}, \quad (6)$$

where $\boldsymbol{\epsilon}_i \stackrel{ind}{\sim} N(0, \boldsymbol{\Sigma}_i)$. The 2×2 matrices \mathbf{R}_i and $\boldsymbol{\Sigma}_i = \mathbf{Q}_i^{-1}$ are respectively defined as:

$$\mathbf{R}_i = \begin{bmatrix} \cos(\alpha d_i) & \frac{1}{\alpha} \sin(\alpha d_i) \\ -\alpha \sin(\alpha d_i) & \cos(\alpha d_i) \end{bmatrix}, \quad \boldsymbol{\Sigma}_i = \sigma^2 \begin{bmatrix} \frac{1}{\alpha^2} \left(\frac{d_i}{2} - \frac{\sin(2\alpha d_i)}{4\alpha} \right) & \frac{\sin^2(\alpha d_i)}{2\alpha^2} \\ \frac{\sin^2(\alpha d_i)}{2\alpha^2} & \frac{2\alpha d_i + \sin(2\alpha d_i)}{4\alpha} \end{bmatrix}. \quad (7)$$

Proof. To show the above Markov representation, note that the value of $g(s_{i+1})$ given $g(s_i)$ can be written similarly as (Särkkä and Solin, 2019):

$$\mathbf{g}_{aug}(s_{i+1}) = \exp(\mathbf{F}d_{i+1})\mathbf{g}_{aug}(s_i) + \int_{s_i}^{s_{i+1}} \exp(\mathbf{F}(s_{i+1} - \tau))\mathbf{J}W(\tau)d\tau.$$

Recall that $\mathbf{F}^{2k} = (-\alpha^2)^k \mathbf{I}$ and $\mathbf{F}^{2k+1} = (-\alpha^2)^k \mathbf{F}$, then apply the Taylor series expansion for both components in the integral above. It then can be rewritten as:

$$\begin{aligned} \mathbf{g}_{aug}(s_{i+1}) &= \exp(\mathbf{F}d_{i+1})\mathbf{g}_{aug}(s_i) + \int_{s_i}^{s_{i+1}} \exp(\mathbf{F}(s_{i+1} - \tau))\mathbf{J}W(\tau)d\tau \\ &= \mathbf{R}_{i+1}\mathbf{g}_{aug}(s_i) + \int_{s_i}^{s_{i+1}} \begin{bmatrix} \frac{1}{\alpha} \sin(\alpha(s_{i+1} - \tau)) \\ \cos(\alpha(s_{i+1} - \tau)) \end{bmatrix} \sigma W(\tau)d\tau \\ &:= \mathbf{R}_{i+1}\mathbf{g}_{aug}(s_i) + \boldsymbol{\epsilon}_{i+1}. \end{aligned} \quad (8)$$

Note that since each $\boldsymbol{\epsilon}_{i+1}$ involves integration at disjoint intervals, their independence follows from the property of Gaussian white noise (Harvey, 1990). To check its covariance matrix $\boldsymbol{\Sigma}_{i+1}$, note that:

$$\begin{aligned} \boldsymbol{\Sigma}_{i+1} &= \sigma^2 \begin{bmatrix} \int_{s_i}^{s_{i+1}} \frac{1}{\alpha^2} \sin^2(\alpha(s_{i+1} - \tau))d\tau & \frac{1}{\alpha} \int_{s_i}^{s_{i+1}} \sin(\alpha(s_{i+1} - \tau)) \cos(\alpha(s_{i+1} - \tau))d\tau \\ \frac{1}{\alpha} \int_{s_i}^{s_{i+1}} \sin(\alpha(s_{i+1} - \tau)) \cos(\alpha(s_{i+1} - \tau))d\tau & \int_{s_i}^{s_{i+1}} \cos^2(\alpha(s_{i+1} - \tau))d\tau \end{bmatrix} \\ &= \sigma^2 \begin{bmatrix} \frac{1}{\alpha^2} \left(\frac{d_{i+1}}{2} - \frac{\sin(2\alpha d_{i+1})}{4\alpha} \right) & \frac{\sin^2(\alpha d_{i+1})}{2\alpha^2} \\ \frac{\sin^2(\alpha d_{i+1})}{2\alpha^2} & \frac{2\alpha d_{i+1} + \sin(2\alpha d_{i+1})}{4\alpha} \end{bmatrix}, \end{aligned} \quad (9)$$

which completes the proof. □

S4. Details of the Finite Element Method

The Finite Element Method (FEM) used to construct the finite-dimensional approximation can be understood as the following procedures.

Given the stochastic differential equation (SDE) that defines the (standard) sGP model:

$$LW(x) = \xi(x),$$

where $L = \alpha^2 + \frac{d^2}{dx^2}$ is a linear differential operator and $\xi(x)$ is the standard Gaussian white noise process. Let $\Omega \subset \mathbb{R}^+$ denotes a bounded interval of interest. Let $\mathbb{B}_k := \{\psi_i, i \in [k]\}$ denote the set of k pre-specified basis functions, and let $\mathbb{T}_q := \{\phi_i, i \in [q]\}$ denote the set of q pre-specified test functions. We consider finite dimensional approximation with form $\widetilde{W}_k(\cdot) = \sum_{i=1}^k w_i \psi_i(\cdot)$. The weights $\mathbf{w} := [w_1, \dots, w_k]^T \in \mathbb{R}^k$ is a set of random weights to be determined.

In our FEM construction, we used the sB-splines defined over Ω as the basis functions, and chose the test functions by $\mathbb{T}_k := \{\phi_i = L\psi_i, i \in [k]\}$, which is called a least squares approximation in Lindgren et al. (2011). The distribution of the unknown weight vector can be found by fulfilling the weak formulation at the test function spaces \mathbb{T}_k , such that

$$\langle L\widetilde{W}_k(x), \phi_i(x) \rangle \stackrel{d}{=} \langle \xi(x), \phi_i(x) \rangle, \quad (10)$$

for any test function $\phi_i \in \mathbb{T}_k$. This equation can also be vectorized as:

$$\langle L\widetilde{W}_k(x), \phi_i(x) \rangle_{i=1}^k = H\mathbf{w},$$

where the ij component of the $k \times k$ H matrix can be computed as $H_{ij} = \langle L\psi_j(x), L\psi_i(x) \rangle_{i=1}^k$.

The inner product on the right $\langle \xi(x), \phi_i(x) \rangle_{i=1}^k$ will have Gaussian distribution with zero mean vector and covariance matrix H by properties of Gaussian white noise (Harvey, 1990). Therefore, the basis coefficients \mathbf{w} will be multivariate Gaussian with zero mean and covari-

ance $H^{-1}HH^{-1} = H^{-1}$. Each element of the matrix H can be written as:

$$\begin{aligned}
H_{ij} &= \langle L\psi_j, L\psi_i \rangle \\
&= \langle a^2\psi_j + \frac{d^2\psi_j}{dx^2}, a^2\psi_i + \frac{d^2\psi_i}{dx^2} \rangle \\
&= a^4\langle \psi_j, \psi_i \rangle + a^2\langle \frac{d^2\psi_j}{dx^2}, \psi_i \rangle + a^2\langle \psi_j, \frac{d^2\psi_i}{dx^2} \rangle + \langle \frac{d^2\psi_j}{dx^2}, \frac{d^2\psi_i}{dx^2} \rangle,
\end{aligned} \tag{11}$$

hence $H = a^4G + C + a^2M$ with $G_{ij} = \langle \psi_i, \psi_j \rangle$, $C_{ij} = \langle \frac{d^2\psi_i}{dx^2}, \frac{d^2\psi_j}{dx^2} \rangle$ and $M_{ij} = \langle \psi_i, \frac{d^2\psi_j}{dx^2} \rangle + \langle \frac{d^2\psi_i}{dx^2}, \psi_j \rangle$ for each element of the matrices.

S5. Proof of the Convergence Result

Theorem (Covariance Convergence of B-spline Approximation). *Assume \mathbb{B}_k is a set of k cubic B-splines constructed with equally spaced knots over Ω , and \tilde{g}_k denotes the corresponding FEM approximation for $sGP_\alpha(\sigma)$, then for any $x_1, x_2 \in \Omega$:*

$$|\mathcal{C}_k(x_1, x_2) - \mathcal{C}(x_1, x_2)| = O(1/k),$$

where $\mathcal{C}(x_1, x_2)$ is the covariance in Proposition S1 and $\mathcal{C}_k(x_1, x_2) = \text{Cov}[\tilde{g}_k(x_1), \tilde{g}_k(x_2)]$.

Proof. The proof of this theorem starts with a similar strategy as in Lindgren et al. (2011). Without the loss of generality, we assume the variance parameter of the sGP $\sigma = 1$, $\Omega = [0, 1]$ and the initial conditions of the sGP are zero. We denote the 3rd order Sobolev space as $H^3(\Omega) = \{f \in \mathcal{L}^2(\Omega) : D^q f \in \mathcal{L}^2(\Omega) \ \forall |q| \leq 3\}$ and the constrained Sobolev space \mathcal{H} as:

$$\mathcal{H} = \{f \in H^3(\Omega) : f(0) = f'(0) = 0\} \subset H^3(\Omega).$$

Since $Lf = 0$ implies $f \in \text{span}\{\cos(\alpha x), \sin(\alpha x)\}$, it is clear that

$$\langle f, h \rangle_{\mathcal{H}} := \langle Lf, Lh \rangle_{\Omega} = \int_{\Omega} Lf(x) Lh(x) dx \quad (12)$$

defines an inner product for $f, h \in \mathcal{H}$.

Define $\mathcal{H}_k = \text{span}\{\mathbb{B}_k\}$, and note $\mathcal{H}_k \subset \mathcal{H}$ by our construction of the B-spline basis. Since \mathcal{H}_k is a finite-dimensional subspace, for each $f(x) \in \mathcal{H}$ there exists a unique projection $\tilde{f}(x) = \sum_{i=1}^k w_i \psi_i(x) \in \mathcal{H}_k$ which satisfies:

$$\langle f - \tilde{f}, \tilde{h} \rangle_{\mathcal{H}} = \langle f, \tilde{h} \rangle_{\mathcal{H}} - \langle \tilde{f}, \tilde{h} \rangle_{\mathcal{H}} = 0, \ \forall \tilde{h} \in \mathcal{H}_k. \quad (13)$$

Based on the property of Gaussian white noise (Harvey, 1990), for any $f, h \in \mathcal{H}$ we have:

$$\text{Cov}[\langle \xi, Lf \rangle_\Omega, \langle \xi, Lh \rangle_\Omega] = \langle Lf, Lh \rangle_\Omega = \langle f, h \rangle_{\mathcal{H}}. \quad (14)$$

Since the FEM approximation $\tilde{g}_k \in \mathcal{H}_k$, we know

$$\langle \tilde{g}_k, f \rangle_{\mathcal{H}} = \langle \tilde{g}_k, f - \tilde{f} + \tilde{f} \rangle_{\mathcal{H}} = \langle \tilde{g}_k, \tilde{f} \rangle_{\mathcal{H}} + \langle \tilde{g}_k, f - \tilde{f} \rangle_{\mathcal{H}} = \langle \tilde{g}_k, \tilde{f} \rangle_{\mathcal{H}},$$

where the last equality follows as \tilde{f} is the projection of f . Using this result and the fact that the FEM approximation \tilde{g}_k is a least square solution, we have

$$\begin{aligned} \text{Cov} \left[\langle \tilde{g}_k, f \rangle_{\mathcal{H}}, \langle \tilde{g}_k, h \rangle_{\mathcal{H}} \right] &= \text{Cov} \left[\langle \tilde{g}_k, \tilde{f} \rangle_{\mathcal{H}}, \langle \tilde{g}_k, \tilde{h} \rangle_{\mathcal{H}} \right] \\ &= \text{Cov} \left[\langle L\tilde{g}_k, L\tilde{f} \rangle_{\Omega}, \langle L\tilde{g}_k, L\tilde{h} \rangle_{\Omega} \right] \\ &= \text{Cov} \left[\langle \xi, L\tilde{f} \rangle_{\Omega}, \langle \xi, L\tilde{h} \rangle_{\Omega} \right] \\ &= \langle L\tilde{f}, L\tilde{h} \rangle_{\Omega} = \langle \tilde{f}, \tilde{h} \rangle_{\mathcal{H}}. \end{aligned} \quad (15)$$

Let $\mathcal{C}_s(x) = \mathcal{C}(s, x)$ denote the covariance function of the sGP defined at any $s \in \Omega$. Based on the previous result in Proposition S1, we know $\mathcal{C}_s(x) \in \mathcal{H}$ and $L\mathcal{C}_s(x) = \frac{1}{\alpha} \sin[\alpha(s - x)^+]$ is the Green function of L . The projection of $\mathcal{C}_s(x)$ into \mathcal{H}_k is denoted as $\tilde{\mathcal{C}}_s(x)$.

Lemma 1. *Given the same setting in the main theorem*

$$\begin{aligned} \mathcal{C}(x_1, x_2) &= \langle C_{x_1}, C_{x_2} \rangle_{\mathcal{H}} \\ \mathcal{C}_k(x_1, x_2) &= \langle \tilde{\mathcal{C}}_{x_1}, \tilde{\mathcal{C}}_{x_2} \rangle_{\mathcal{H}}. \end{aligned} \quad (16)$$

Proof. The first part directly follows from the proof in Proposition S1. The second part can be proved using the fact that $L\mathcal{C}_{x_1}(x)$ is the Green function, which implies

$$\langle \psi_i, \tilde{\mathcal{C}}_{x_1} \rangle_{\mathcal{H}} = \langle \psi_i, \mathcal{C}_{x_1} \rangle_{\mathcal{H}} = \langle L\psi_i, L\mathcal{C}_{x_1} \rangle_{\Omega} = \psi_i(x_1),$$

for each $\psi_i \in \mathbb{B}_k$. The detailed proof proceeds as follow.

By construction of the B-spline approximation,

$$\begin{aligned}
\mathcal{C}_k(x_1, x_2) &= \text{Cov} \left[\sum_i w_i \psi_i(x_1), \sum_i w_i \psi_i(x_2) \right] \\
&= \text{Cov} \left[\mathbf{\Phi}(x_1)^T \mathbf{w}, \mathbf{\Phi}(x_2)^T \mathbf{w} \right] \\
&= \mathbf{\Phi}(x_1)^T \Sigma_{\mathbf{w}} \mathbf{\Phi}(x_2) \\
&= \boldsymbol{\gamma}_{x_1}^T \Sigma_{\mathbf{w}}^{-1} \boldsymbol{\gamma}_{x_2},
\end{aligned} \tag{17}$$

where $\mathbf{\Phi}(x) = [\psi_1(x), \dots, \psi_k(x)]^T$, and $\Sigma_{\mathbf{w}}$ is defined in Section 3.2, and $\boldsymbol{\gamma}_{x_1} = \Sigma_{\mathbf{w}} \mathbf{\Phi}(x_1)$ and $\boldsymbol{\gamma}_{x_2} = \Sigma_{\mathbf{w}} \mathbf{\Phi}(x_2)$.

Since $\tilde{\mathcal{C}}_{x_1}(x)$ is the projection of $\mathcal{C}_{x_1}(x)$ to \mathcal{H}_k , $\tilde{\mathcal{C}}_{x_1}(x) = \sum_i w_{x_1, i} \psi_i(x)$ for some weights $\mathbf{w}_{x_1} = [w_{x_1, 1}, \dots, w_{x_1, k}]^T$. The same argument can be used for $\tilde{\mathcal{C}}_{x_2}$. Therefore

$$\begin{aligned}
\text{Cov} \left[\left\langle \tilde{g}_k, \tilde{\mathcal{C}}_{x_1} \right\rangle_{\mathcal{H}}, \left\langle \tilde{g}_k, \tilde{\mathcal{C}}_{x_2} \right\rangle_{\mathcal{H}} \right] &= \left\langle \tilde{\mathcal{C}}_{x_1}, \tilde{\mathcal{C}}_{x_2} \right\rangle_{\mathcal{H}} \\
&= \left\langle \mathbf{\Phi}(x)^T \mathbf{w}_{x_1}, \mathbf{\Phi}(x)^T \mathbf{w}_{x_2} \right\rangle_{\mathcal{H}} \\
&= \left\langle L \mathbf{\Phi}(x)^T \mathbf{w}_{x_1}, L \mathbf{\Phi}(x)^T \mathbf{w}_{x_2} \right\rangle_{\Omega} \\
&= \mathbf{w}_{x_1}^T \Sigma_{\mathbf{w}}^{-1} \mathbf{w}_{x_2},
\end{aligned} \tag{18}$$

since $[\Sigma_{\mathbf{w}}^{-1}]_{ij} = \langle L \psi_i, L \psi_j \rangle_{\Omega}$, hence it only remains to show $\boldsymbol{\gamma}_{x_1} = \mathbf{w}_{x_1}$. Since

$$\begin{aligned}
\left\langle \tilde{\mathcal{C}}_{x_1}, \psi_i \right\rangle_{\mathcal{H}} &= \langle \mathcal{C}_{x_1}, \psi_i \rangle_{\mathcal{H}} = \psi_i(x_1), \\
\left\langle \tilde{\mathcal{C}}_{x_1}, \psi_i \right\rangle_{\mathcal{H}} &= \left\langle \mathbf{\Phi}(x)^T \mathbf{w}_{x_1}, \psi_i \right\rangle_{\mathcal{H}} = \sum_{j=1}^k w_{x_1, j} \langle L \psi_j, L \psi_i \rangle_{\Omega},
\end{aligned} \tag{19}$$

for each $\psi_i \in \mathbb{B}_k$, we get $\Sigma_{\mathbf{w}}^{-1} \mathbf{w}_{x_1} = \mathbf{\Phi}(x_1)$. This lemma is hence proved. \square

Using the above result and Lemma 1, it suffices to prove that

$$\left| \left\langle \tilde{\mathcal{C}}_{x_1}, \tilde{\mathcal{C}}_{x_2} \right\rangle_{\mathcal{H}} - \langle \mathcal{C}_{x_1}, \mathcal{C}_{x_2} \rangle_{\mathcal{H}} \right| = O(1/k). \tag{20}$$

For this step, we will use the following lemma on the spline approximation:

Lemma 2. *Given the same setting in the main theorem, define the norm $\|f\|_{\mathcal{H}} = \langle f, f \rangle_{\mathcal{H}}^{1/2}$ for $f \in \mathcal{H}$ then*

$$\|\mathcal{C}_s - \tilde{\mathcal{C}}_s\|_{\mathcal{H}} = O(1/k), \quad (21)$$

for each $s \in \Omega$.

Proof. The proof of this lemma mostly follows from the result in Schultz (1969). First, since $D^2\mathcal{C}_s(x)$ is a continuous function on Ω and differentiable everywhere except at $x = s$, $\mathcal{C}_s(x)$ has weak derivatives up to order 3. As the derivative of $D^2\mathcal{C}_s(x)$ is bounded and continuous for $x < s$ and $x > s$, we can conclude $\mathcal{C}_s \in H^3(\Omega)$. Given \mathcal{H}_k is a spline space with degree 3 and mesh size $1/k$ and $\mathcal{C}_s \in H^3(\Omega)$, by theorem 3.3 in Schultz (1969) we have

$$\|D^q(\mathcal{C}_s - \tilde{\mathcal{C}}_s)\|_{\mathcal{L}^2} \leq c_q \left(\frac{1}{k}\right)^{3-q} \quad (22)$$

for each $0 \leq q \leq 2$, where c_q is a constant that only depends on $\|\mathcal{C}_s\|_{H^3(\Omega)}$. Note that

$$\begin{aligned} \|\mathcal{C}_s - \tilde{\mathcal{C}}_s\|_{\mathcal{H}}^2 &= \alpha^4 \|\mathcal{C}_s - \tilde{\mathcal{C}}_s\|_{\mathcal{L}^2}^2 + \|D^2(\mathcal{C}_s - \tilde{\mathcal{C}}_s)\|_{\mathcal{L}^2}^2 - 2\alpha \left\langle (\mathcal{C}_s - \tilde{\mathcal{C}}_s), D^2(\mathcal{C}_s - \tilde{\mathcal{C}}_s) \right\rangle_{\Omega} \\ &\leq \alpha^4 \|\mathcal{C}_s - \tilde{\mathcal{C}}_s\|_{\mathcal{L}^2}^2 + \|D^2(\mathcal{C}_s - \tilde{\mathcal{C}}_s)\|_{\mathcal{L}^2}^2 + 2\alpha \|D^2(\mathcal{C}_s - \tilde{\mathcal{C}}_s)\|_{\mathcal{L}^2} \|\mathcal{C}_s - \tilde{\mathcal{C}}_s\|_{\mathcal{L}^2} \end{aligned} \quad (23)$$

where the second equality holds by Cauchy-Schwarz Inequality. The lemma hence proved. \square

Using Lemma 2, Eq. (20) can be proved with an application of Triangle Inequality followed with a use of Cauchy-Schwarz Inequality and the fact that the sequence $\|\tilde{\mathcal{C}}_{x_2}\|_{\mathcal{H}}$ is bounded,

$$\begin{aligned} \left| \left\langle \tilde{\mathcal{C}}_{x_1}, \tilde{\mathcal{C}}_{x_2} \right\rangle_{\mathcal{H}} - \langle \mathcal{C}_{x_1}, \mathcal{C}_{x_2} \rangle_{\mathcal{H}} \right| &= \left| \left\langle \tilde{\mathcal{C}}_{x_1} - \mathcal{C}_{x_1}, \tilde{\mathcal{C}}_{x_2} \right\rangle_{\mathcal{H}} - \left\langle \mathcal{C}_{x_1}, \mathcal{C}_{x_2} - \tilde{\mathcal{C}}_{x_2} \right\rangle_{\mathcal{H}} \right| \\ &\leq \|\tilde{\mathcal{C}}_{x_1} - \mathcal{C}_{x_1}\|_{\mathcal{H}} \|\tilde{\mathcal{C}}_{x_2}\|_{\mathcal{H}} + \|\tilde{\mathcal{C}}_{x_2} - \mathcal{C}_{x_2}\|_{\mathcal{H}} \|\mathcal{C}_{x_1}\|_{\mathcal{H}} \\ &\leq c/k, \end{aligned} \quad (24)$$

where c is some constant independent of k . The theorem is hence proved. □

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